

Online Supplementary Document

Bayesian Bandwidth Estimation in Nonparametric Time–Varying Coefficient Models¹

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Appendix C: Proofs of Theorems 1 and 2

Observe that

$$\begin{aligned}\widehat{v}_t &= \widehat{u}_t - \alpha \widehat{u}_{t-1} = y_t - x_t^\top \widehat{\beta}(\tau_t; h) - \alpha(y_{t-1} - x_{t-1}^\top \widehat{\beta}(\tau_{t-1}; h)) \\ &= u_t + x_t^\top (\beta(\tau_t) - \widehat{\beta}(\tau_t; h)) - \alpha(u_{t-1} + x_{t-1}^\top (\beta(\tau_{t-1}) - \widehat{\beta}(\tau_{t-1}; h))) \\ &= u_t - \alpha u_{t-1} + \Delta(\tau_t; h) - \alpha \Delta(\tau_{t-1}; h) \\ &= v_t + \Delta(\tau_t; h) - \alpha \Delta(\tau_{t-1}; h) = v_t + \Gamma(\tau_t; \alpha, h),\end{aligned}$$

where $\Delta(\tau_t; h) = x_t^\top (\beta(\tau_t) - \widehat{\beta}(\tau_t; h))$ and $\Gamma(\tau_t; \alpha, h) = \Delta(\tau_t; h) - \alpha \Delta(\tau_{t-1}; h)$.

Therefore, we can further show

$$l_n(\lambda, \theta) = \log L_n(\lambda, \theta) = \sum_{t=1}^n \log f(\widehat{v}_t; \theta)$$

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$$\begin{aligned}
&= \sum_{t=1}^n \log f(v_t + \Gamma(\tau_t; \alpha, h); \theta) \\
&= \sum_{t=1}^n \log [(f(v_t; \theta) + f^{(1)}(v_t; \theta)\Gamma(\tau_t; \alpha, h))(1 + o_P(1))] \\
&= \sum_{t=1}^n \log \left[f(v_t; \theta) \left(1 + \frac{f^{(1)}(v_t; \theta)}{f(v_t; \theta)} \Gamma(\tau_t; \alpha, h) \right) (1 + o_P(1)) \right] \\
&= \sum_{t=1}^n \log f(v_t; \theta) + \sum_{t=1}^n \log (1 + \gamma(v_t; \theta)\Gamma(\tau_t; \alpha, h)) + o_P(1) \\
&= \sum_{t=1}^n \log f(v_t; \theta) + \sum_{t=1}^n \gamma(v_t; \theta)\Gamma(\tau_t; \alpha, h) (1 + o_P(1)),
\end{aligned} \tag{1}$$

where $\gamma(v_t; \theta) = \frac{f^{(1)}(v_t; \theta)}{f(v_t; \theta)}$.

Thus, we have

$$\begin{aligned}
L_n(\lambda, \theta) &= e^{\sum_{t=1}^n \log f(v_t; \theta)} e^{\sum_{t=1}^n \gamma(v_t; \theta)\Gamma(\tau_t; \alpha, h)(1 + o_P(1))} \\
&= (1 + o_P(1)) \prod_{t=1}^n f(v_t; \theta) \left(\sum_{t=1}^n \gamma(v_t; \theta)\Gamma(\tau_t; \alpha, h) + 1 \right) \\
&= (1 + o_P(1)) G_n(\theta) \left(\sum_{t=1}^n \gamma(v_t; \theta)\Gamma(\tau_t; \alpha, h) + 1 \right),
\end{aligned} \tag{2}$$

where $G_n(\theta) = \prod_{t=1}^n f(v_t; \theta)$.

It is easy to see that

$$L_n(\lambda, \theta) - G_n(\theta) = G_n(\theta) \sum_{t=1}^n \gamma(v_t; \theta)\Gamma(\tau_t; \alpha, h).$$

As $\Gamma(\tau_t; \alpha, h) = \Delta(\tau_t; h) - \alpha\Delta(\tau_{t-1}; h)$, we have

$$\sum_{t=1}^n \gamma(v_t; \theta)\Gamma(\tau_t; \alpha, h) = \sum_{t=1}^n \gamma(v_t; \theta)\Delta(\tau_t; h) - \alpha \sum_{t=1}^n \gamma(v_t; \theta)\Delta(\tau_{t-1}; h).$$

We investigate $\Delta(\tau_t; h)$ as follows:

$$\begin{aligned}
\Delta(\tau_t; h) &= x_t^\top (\beta(\tau_t) - \widehat{\beta}(\tau_t; h)) = x_t^\top (\beta(\tau_t) - [p_n(\tau_t; h)]^{-1} q_n(\tau_t; h)) \\
&= x_t^\top \left\{ [p_n(\tau_t; h)]^{-1} [p_n(\tau_t; h)] \beta(\tau_t) - [p_n(\tau_t; h)]^{-1} \frac{1}{nh} \sum_{s=1}^n x_s (x_s^\top \beta(\tau_s) + u_s) K\left(\frac{\tau_s - \tau_t}{h}\right) \right\} \\
&= -x_t^\top [p_n(\tau_t; h)]^{-1} \frac{1}{nh} \sum_{s=1}^n x_s u_s K\left(\frac{\tau_s - \tau_t}{h}\right) \\
&\quad + x_t^\top [p_n(\tau_t; h)]^{-1} \frac{1}{nh} \sum_{s=1}^n x_s x_s^\top (\beta(\tau_t) - \beta(\tau_s)) K\left(\frac{\tau_s - \tau_t}{h}\right) = -\Delta_1(\tau_t; h) + \Delta_2(\tau_t; h),
\end{aligned}$$

in which $p_n(\tau_t; h) = \frac{1}{nh} \sum_{s=1}^n x_s x_s^\top K\left(\frac{\tau_s - \tau_t}{h}\right)$, $q_n(\tau_t; h) = \frac{1}{nh} \sum_{s=1}^n x_s y_s K\left(\frac{\tau_s - \tau_t}{h}\right)$,

$\Delta_1(\tau_t; h) = x_t^\top [p_n(\tau_t; h)]^{-1} \frac{1}{nh} \sum_{s=1}^n x_s u_s K\left(\frac{\tau_s - \tau_t}{h}\right)$, and

$\Delta_2(\tau_t; h) = x_t^\top [p_n(\tau_t; h)]^{-1} \frac{1}{nh} \sum_{s=1}^n x_s x_s^\top (\beta(\tau_t) - \beta(\tau_s)) K\left(\frac{\tau_s - \tau_t}{h}\right)$.

Therefore, we have

$$\sum_{t=1}^n \gamma(v_t; \theta) \Delta(\tau_t; h) = - \sum_{t=1}^n \gamma(v_t; \theta) \Delta_1(\tau_t; h) + \sum_{t=1}^n \gamma(v_t; \theta) \Delta_2(\tau_t; h). \quad (3)$$

Note that $p_n(\tau; h)$ can be expressed as

$$\begin{aligned} p_n(\tau; h) &= \frac{1}{nh} \sum_{s=1}^n x_s x_s^\top K\left(\frac{\tau_s - \tau}{h}\right) = (1 + o(1)) \mathbb{E}[x_1 x_1^\top] \frac{1}{nh} \sum_{s=1}^n K\left(\frac{\tau_s - \tau}{h}\right) \\ &= (1 + o(1)) \mathbb{E}[x_1 x_1^\top] \int \frac{1}{h} K\left(\frac{u - \tau}{h}\right) du = (1 + o(1)) \mathbb{E}[x_1 x_1^\top] \int_{-\frac{\tau}{h}}^{\frac{1-\tau}{h}} k(v) dv. \end{aligned}$$

As we assume that $h = a_n \lambda$ with $a_n \rightarrow 0$ as $n \rightarrow \infty$ and λ being a continuous random variable, so $h \rightarrow 0$ as $n \rightarrow \infty$. Therefore, $\int_{-\frac{\tau}{h}}^{\frac{1-\tau}{h}} k(v) dv = 1 + o(1)$ and

$$p_n(\tau; h) = (1 + o(1)) \mathbb{E}[x_1 x_1^\top] = \Sigma_x (1 + o(1)).$$

Therefore, it follows that

$$\begin{aligned} \sum_{t=1}^n \gamma(v_t; \theta) \Delta_1(\tau_t; h) &= \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top [p_n(\tau_t; h)]^{-1} \frac{1}{nh} \sum_{s=1}^n x_s u_s K\left(\frac{\tau_s - \tau_t}{h}\right) \\ &= \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} \frac{1}{nh} \sum_{s=1}^n x_s u_s K\left(\frac{\tau_s - \tau_t}{h}\right) \\ &= \frac{1}{nh} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} \sum_{s=1}^n x_s u_s K\left(\frac{\tau_s - \tau_t}{h}\right) \\ &= \frac{K(0)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t u_t + \frac{1}{nh} \sum_{s=1}^n \sum_{t=1, t \neq s}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s u_s K\left(\frac{\tau_s - \tau_t}{h}\right) \\ &= I_{nt}(h) + J_{nt}(h), \end{aligned}$$

where $I_{nt}(h) = \frac{K(0)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t u_t$ and $J_{nt}(h) = \frac{1}{nh} \sum_{s=1}^n \sum_{t=1, t \neq s}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s u_s K\left(\frac{\tau_s - \tau_t}{h}\right)$.

Since the error term u_t follows an AR(1) process, we can further express $I_{nt}(h)$ and $J_{nt}(h)$ as follows:

$$I_{nt}(h) = \frac{K(0)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t u_t = \frac{K(0)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t (\alpha u_{t-1} + v_t)$$

$$\begin{aligned}
&= \alpha \frac{K(0)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t u_{t-1} + \frac{K(0)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t v_t, \\
J_{nt}(h) &= \frac{1}{nh} \sum_{s=1}^n \sum_{t=1, t \neq s}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s u_s K\left(\frac{\tau_s - \tau_t}{h}\right) \\
&= \frac{1}{nh} \sum_{s=1}^n \sum_{t=1, t \neq s}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s (\alpha u_{s-1} + v_s) K\left(\frac{\tau_s - \tau_t}{h}\right) \\
&= \alpha \frac{1}{nh} \sum_{s=1}^n \sum_{t=1, t \neq s}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s u_{s-1} K\left(\frac{\tau_s - \tau_t}{h}\right) \\
&\quad + \frac{1}{nh} \sum_{s=1}^n \sum_{t=1, t \neq s}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s v_s K\left(\frac{\tau_s - \tau_t}{h}\right).
\end{aligned}$$

Similarly, we can get

$$\alpha \sum_{t=1}^n \gamma(v_t; \theta) \Delta(\tau_{t-1}; h) = -\alpha \sum_{t=1}^n \gamma(v_t; \theta) \Delta_1(\tau_{t-1}; h) + \alpha \sum_{t=1}^n \gamma(v_t; \theta) \Delta_2(\tau_{t-1}; h), \quad (4)$$

where $\sum_{t=1}^n \gamma(v_t; \theta) \Delta_1(\tau_{t-1}; h) = I_{n(t-1)}(h) + J_{n(t-1)}(h)$, $I_{n(t-1)}(h) = \frac{K(0)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) x_{t-1}^\top \Sigma_x^{-1} x_{t-1} u_{t-1}$ and $J_{n(t-1)}(h) = \frac{1}{nh} \sum_{s=1}^n \sum_{t=1, t \neq s}^n \gamma(v_t; \theta) x_{t-1}^\top \Sigma_x^{-1} x_s u_s K\left(\frac{\tau_s - \tau_{t-1}}{h}\right)$.

Observe that

$$\begin{aligned}
&\sum_{t=1}^n \gamma(v_t; \theta) \Delta_1(\tau_t; h) - \alpha \sum_{t=1}^n \gamma(v_t; \theta) \Delta_1(\tau_{t-1}; h) \\
&= \frac{K(0)}{nh} \sum_{t=1}^n v_t \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t + \frac{1}{nh} \sum_{s=1}^n \sum_{t=1, t \neq s}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s v_s K\left(\frac{\tau_s - \tau_t}{h}\right) + o_P(1) \\
&= Q_{1n} + Q_{2n} + o_P(1),
\end{aligned}$$

where $Q_{1n} = \frac{K(0)}{nh} \sum_{t=1}^n v_t \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t$ and $Q_{2n} = \frac{1}{nh} \sum_{s=1}^n \sum_{t=1, t \neq s}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s v_s K\left(\frac{\tau_s - \tau_t}{h}\right)$.

As $n \rightarrow \infty$, the law of large numbers implies

$$hQ_{1n} = K(0) \frac{1}{n} \sum_{t=1}^n v_t \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t \xrightarrow{P} K(0) \cdot \mathbb{E}_\theta[\gamma(v_1; \theta) v_1] \cdot \mathbb{E}[w_1] \neq 0,$$

where $w_t = x_t^\top \Sigma_x^{-1} x_t$. So $Q_{1n} = O_P(1/h)$.

We now show that Q_{2n} is a higher-order term than Q_{1n} . Because $\mathbb{E}_\theta[v_1] = \mathbb{E}_\theta[\gamma(v_1; \theta)] = 0$, we have

$$\begin{aligned}
\mathbb{E}_{\theta, \lambda} [Q_{2n}^2] &= \frac{1}{n^2 h^2} \mathbb{E}_{\theta, \lambda} \left[\sum_{t=1}^n \left(\sum_{s=1, s \neq t}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s v_s K\left(\frac{\tau_s - \tau_t}{h}\right) \right)^2 \right] \\
&= \frac{1}{n^2 h^2} \sum_{t=1}^n \left(\sum_{s=1, s \neq t}^n \mathbb{E}_\theta [\gamma^2(v_t; \theta)] \mathbb{E}_\theta [v_s^2] \cdot \mathbb{E}[w_s w_t] K^2\left(\frac{\tau_s - \tau_t}{h}\right) \right)
\end{aligned}$$

$$\begin{aligned}
&\leq \frac{C_{vw}}{n^2 h^2} \sum_{t=1}^n \sum_{s=1, t \neq s}^n K^2\left(\frac{\tau_s - \tau_t}{h}\right) = \frac{C_{vw}}{nh^2} \sum_{t=1}^n \frac{1}{n} \sum_{s=1, t \neq s}^n K^2\left(\frac{\tau_s - \tau_t}{h}\right) \\
&= \frac{C_{vw}(1 + o(1))}{nh^2} \sum_{t=1}^n \int_0^1 K^2\left(\frac{u - \tau_t}{h}\right) du = \frac{C_{vw}}{nh} \sum_{t=1}^n \int_{-\tau_t/h}^{1-\tau_t/h} K^2(v) dv \\
&= \frac{C_{vw}}{h} \frac{1}{n} \sum_{t=1}^n m(\tau_t; h) = \frac{C_{vw}(1 + o(1))}{h} \int K^2(u) du,
\end{aligned}$$

where $C_{vw} > 0$ is some constant and $m(\tau; h) = \int_{-\tau/h}^{1-\tau/h} K^2(u) du = (1 + o(1)) \int K^2(u) du$. So we have $Q_{2n} = O_P(1/\sqrt{h})$. This deduces that $Q_{1n} + Q_{2n} = O_P(1/h) \left(1 + O_P(\sqrt{h})\right)$, which implies that Q_{1n} is the leading term.

Meanwhile, we have

$$\begin{aligned}
\sum_{t=1}^n \gamma(v_t; \theta) \Delta_2(\tau_t; h) &= \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top [p_n(\tau_t; h)]^{-1} \frac{1}{nh} \sum_{s=1}^n x_s x_s^\top (\beta(\tau_t) - \beta(\tau_s)) K\left(\frac{\tau_s - \tau_t}{h}\right) \\
&= \frac{1 + o(1)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} \sum_{s=1}^n x_s x_s^\top K\left(\frac{\tau_s - \tau_t}{h}\right) (\beta(\tau_t) - \beta(\tau_s)) \\
&= \frac{1 + o(1)}{nh} \sum_{t=1}^n \sum_{s=1, s \neq t}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top K\left(\frac{\tau_s - \tau_t}{h}\right) (\beta^{(1)}(\tau_t)(\tau_t - \tau_s) + 1/2 \beta^{(2)}(\tau_t)(\tau_t - \tau_s)^2) \\
&= \frac{1 + o(1)}{n} \sum_{t=1}^n \sum_{s=1, s \neq t}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top K_1\left(\frac{\tau_s - \tau_t}{h}\right) \beta^{(1)}(\tau_t) \\
&\quad + \frac{(1 + o(1))h}{2n} \sum_{t=1}^n \sum_{s=1, s \neq t}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top K_2\left(\frac{\tau_s - \tau_t}{h}\right) \beta^{(2)}(\tau_t) \\
&\equiv (1 + o(1)) Q_{3n} + (1 + o(1)) Q_{4n}. \tag{5}
\end{aligned}$$

We then show that $\mathbb{E}_{\theta, \lambda}[Q_{jn}] \rightarrow 0$ and $\mathbb{E}_{\theta, \lambda} (Q_{jn} - \mathbb{E}_{\theta, \lambda}[Q_{jn}])^2 \rightarrow 0$ for $j = 3, 4$. Obviously, we have

$$\begin{aligned}
\mathbb{E}_{\theta, \lambda}[Q_{3n}] &= \frac{1}{n} \sum_{t=1}^n \sum_{s=1, s \neq t}^n \mathbb{E}_\theta [\gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top] K_1\left(\frac{\tau_s - \tau_t}{h}\right) \beta^{(1)}(\tau_t), \\
\mathbb{E}_{\theta, \lambda}[Q_{4n}] &= \frac{h}{2n} \sum_{t=1}^n \sum_{s=1, s \neq t}^n \mathbb{E}_\theta [\gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top] K_2\left(\frac{\tau_s - \tau_t}{h}\right) \beta^{(2)}(\tau_t),
\end{aligned}$$

where $K_1(u) = uK(u)$ and $K_2(u) = u^2K(u)$.

Under Assumption 3(ii), by Lemma A.1 of Gao (2007), we obtain³

$$\begin{aligned}
& \|\mathbb{E}_\theta [\gamma(\nu_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top] - \mathbb{E}_\theta [\gamma(\nu_t; \theta) x_t^\top \Sigma_x^{-1}] \mathbb{E}[x_s x_s^\top]\| \\
&= \|\mathbb{E}_\theta [\gamma(\nu_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top] - \mathbb{E}_\theta [\gamma(\nu_t; \theta)] \mathbb{E}[x_t^\top \Sigma_x^{-1}] \mathbb{E}[x_s x_s^\top]\| \\
&= \|\mathbb{E}_\theta [\gamma(\nu_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top]\| \\
&\leq (\mathbb{E}_\theta [\|\gamma(\nu_t; \theta) x_t^\top \Sigma_x^{-1}\|^p])^{1/p} (\mathbb{E}[\|x_s x_s^\top\|^q])^{1/q} \rho^{1-Q}(|s-t|),
\end{aligned}$$

where $Q = \frac{1}{p} + \frac{1}{q} < 1$. Without loss of generosity, we choose $p = q = 4$. Then under Assumptions 1(i) and 3(ii), we have

$$\begin{aligned}
\|\mathbb{E}_\theta [\gamma(\nu_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top]\| &\leq (\mathbb{E}_\theta [\|\gamma(\nu_t; \theta) x_t^\top \Sigma_x^{-1}\|^4])^{1/4} (\mathbb{E}[\|x_s x_s^\top\|^4])^{1/4} \rho^{1/2}(|s-t|) \\
&= (\mathbb{E}_\theta [\gamma^4(\nu_t; \theta)])^{1/4} (\mathbb{E}[\|x_t^\top \Sigma_x^{-1}\|^4])^{1/4} (\mathbb{E}[\|x_s x_s^\top\|^4])^{1/4} \rho^{1/2}(|s-t|) \equiv A_0 \rho^{1/2}(|s-t|),
\end{aligned}$$

where $A_0 = (\mathbb{E}_\theta [\gamma^4(\nu_t; \theta)])^{1/4} (\mathbb{E}[\|x_t^\top \Sigma_x^{-1}\|^4])^{1/4} (\mathbb{E}[\|x_s x_s^\top\|^4])^{1/4}$.

Therefore, we obtain

$$\begin{aligned}
\left| \sum_{t=1}^n \gamma(\nu_t; \theta) \Delta_2(\tau_t; h) \right| &\leq \frac{M}{n} \sum_{t=1}^n \sum_{s=1, s \neq t}^n \|\mathbb{E}_\theta [\gamma(\nu_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top]\| \left| K_1 \left(\frac{\tau_s - \tau_t}{h} \right) \right| \|\beta^{(1)}(\tau_t)\| \\
&\quad + \frac{Mh}{2n} \sum_{t=1}^n \sum_{s=1, s \neq t}^n \|\mathbb{E}_\theta [\gamma(\nu_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top]\| K_2 \left(\frac{\tau_s - \tau_t}{h} \right) \|\beta^{(2)}(\tau_t)\| \\
&\leq \frac{M}{n} \sum_{t=1}^n \sum_{s=1, s \neq t}^n \rho^{1/2}(|s-t|) \left| K_1 \left(\frac{\tau_s - \tau_t}{h} \right) \right| \|\beta^{(1)}(\tau_t)\| \\
&\quad + \frac{Mh}{2n} \sum_{t=1}^n \sum_{s=1, s \neq t}^n \rho^{1/2}(|s-t|) K_2 \left(\frac{\tau_s - \tau_t}{h} \right) \|\beta^{(2)}(\tau_t)\| \\
&= \frac{M}{n} \sum_{t=1}^n \sum_{s=1, s \neq t}^n \rho^{1/2}(n|\tau_s - \tau_t|) \left| K_1 \left(\frac{\tau_s - \tau_t}{h} \right) \right| \|\beta^{(1)}(\tau_t)\| \\
&\quad + \frac{Mh}{2n} \sum_{t=1}^n \sum_{s=1, s \neq t}^n \rho^{1/2}(n|\tau_s - \tau_t|) K_2 \left(\frac{\tau_s - \tau_t}{h} \right) \|\beta^{(2)}(\tau_t)\| \\
&= Mn (1 + o(1)) \int_0^1 \int_0^1 \rho^{1/2}(n|v - u|) \left| K_1 \left(\frac{v - u}{h} \right) \right| \|\beta^{(1)}(u)\| du dv
\end{aligned}$$

³As discussed in the remark about Assumption 1 in Appendix A1, we need only to require $\mathbb{E}_\theta [\gamma(\nu_t; \theta) x_t] = \mathbb{E}_\theta [\gamma(\nu_t; \theta)] \mathbb{E}[x_t] = 0$ in an application of Lemma A.1 of Gao (2007).

Meanwhile, we will have $\mathbb{E}_\theta [\gamma(\nu_t; \theta) x_t^\top \Sigma_x^{-1} x_s x_s^\top] = \mathbb{E}_\theta [\gamma(\nu_t; \theta)] \mathbb{E}[x_t^\top \Sigma_x^{-1} x_s x_s^\top] = 0$ if we assume that $\{\nu_t\}$ and $\{x_s\}$ are independent for all (s, t) . In this case, there is no need to apply Lemma A.1 of Gao (2007) with the choice of (p, q) . As a matter of fact, $\mathbb{E}_{\theta, \lambda} [Q_{jn}] = 0$ and the derivation of $\mathbb{E}_{\theta, \lambda} [Q_{jn}]^2 \rightarrow 0$ for $j = 3, 4$ follows similarly from that of $\mathbb{E}_{\theta, \lambda} [Q_{2n}]^2 \rightarrow 0$.

$$\begin{aligned}
& + \frac{Mnh}{2} (1 + o(1)) \int_0^1 \int_0^1 \rho^{1/2}(n|v-u|) K_2\left(\frac{v-u}{h}\right) \|\beta^{(2)}(u)\| du dv \\
& = nhM \cdot A_{1n} \int_0^1 \|\beta^{(1)}(u)\| du + \frac{Mnh^2}{2} \cdot A_{2n} \int_0^1 \|\beta^{(2)}(u)\| du,
\end{aligned}$$

where $A_{1n} = \int_{-\infty}^{+\infty} \rho^{1/2}(nh|w|) |K_1(w)| dw$, $A_{2n} = \int_{-\infty}^{+\infty} \rho^{1/2}(nh|w|) K_2(w) dw$ and M is a positive constant.

Under Assumption 1(i), we have

$$\begin{aligned}
|A_{1n}| &= \left| \int_{-\infty}^{+\infty} \rho^{1/2}(nh|w|) |K_1(w)| dw \right| \leq \int_{-\infty}^{+\infty} \rho^{1/2}(nh|w|) |w| K(w) dw \\
&\leq c(nh)^{-c_0/2} \int_{-\infty}^{+\infty} |w|^{-c_0/2} |w| K(w) dw \leq c(nh)^{-c_0/2} \int_{-\infty}^{+\infty} |w|^{1-c_0/2} K(w) dw.
\end{aligned}$$

Thus, under Assumptions 1(i)(iii) and 2(i), as $n \rightarrow \infty$, we have

$$nh|A_{1n}| \leq c(nh)^{1-c_0/2} \int_{-\infty}^{+\infty} |w|^{1-c_0/2} K(w) dw = O((nh)^{1-c_0/2}) \rightarrow 0.$$

Similarly, we have

$$\begin{aligned}
|A_{2n}| &= \left| \int_{-\infty}^{+\infty} \rho^{1/2}(nh|w|) K_2(w) dw \right| \leq \int_{-\infty}^{+\infty} \rho^{1/2}(nh|w|) |w|^2 K(w) dw \\
&\leq c(nh)^{-c_0/2} \int_{-\infty}^{+\infty} |w|^{2-c_0/2} K(w) dw.
\end{aligned}$$

Hence, under Assumptions 1(i)(iii) and 2(i), as $n \rightarrow \infty$, we have

$$nh^2|A_{2n}| \leq cn^{1-c_0/2} h^{2-c_0/2} \int_{-\infty}^{+\infty} |w|^{2-c_0/2} K(w) dw = O(n^{1-c_0/2} h^{2-c_0/2}) \rightarrow 0.$$

Thus, under Assumptions 1(i)(iii), 2(i) and 3(i)(ii), it follows that

$$\left| \sum_{t=1}^n \gamma(v_t; \theta) \Delta_2(\tau_t; h) \right| \leq nhM \int_0^1 \|\beta^{(1)}(u)\| du A_{1n} + \frac{nh^2}{2} M \int_0^1 \|\beta^{(2)}(u)\| du A_{2n} = o_P(1).$$

Similarly, we can show that $\sum_{t=1}^n \gamma(v_t; \theta) \Delta_2(\tau_{t-1}; h) = o_P(1)$. Under Assumptions 1(ii) and 1(iii), we therefore have

$$\begin{aligned}
& \int \left| \sum_{t=1}^n \gamma(v_t; \theta) \Delta_2(\tau_t; h) \right| \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta \\
& \leq \int \tilde{c}_1(nh)^{1-c_0/2} \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta + \int \tilde{c}_2 n^{1-c_0/2} h^{2-c_0/2} \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta \\
& = \tilde{c}_1 n^{1-c_0/2} a_n^{1-c_0/2} \int \lambda^{1-c_0/2} \pi_\lambda(\lambda) d\lambda + \tilde{c}_2 n^{1-c_0/2} a_n^{2-c_0/2} \int \lambda^{2-c_0/2} \pi_\lambda(\lambda) d\lambda \rightarrow 0,
\end{aligned}$$

where $|\tilde{c}_1| < \infty$ and $|\tilde{c}_2| < \infty$.

In order to show that $\mathbb{E}_{\theta,\lambda} \left[\left(Q_{jn} - \mathbb{E}_{\theta,\lambda} [Q_{jn}^2] \right)^2 \right] \rightarrow 0$ for $j = 3, 4$, we need only to deal with a leading term of the form

$$\frac{1}{n^2} \mathbb{E}_{\theta,\lambda} \left[\left(\sum_{t=1}^n \sum_{s=1, s \neq t}^n (\gamma(v_t; \theta) x_t^\top \Sigma_x^{-1}) (x_s x_s^\top - \mathbb{E}[x_s x_s^\top]) K_1 \left(\frac{\tau_s - \tau_t}{h} \right) \beta^{(1)}(\tau_t) \right)^2 \right] \rightarrow 0,$$

which follows similarly from the derivation of $\mathbb{E}[Q_{2n}^2] \rightarrow 0$.

Let $w_t = x_t^\top \Sigma_x^{-1} x_t$. To sum up, we have

$$\begin{aligned} \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) &= \sum_{t=1}^n \gamma(v_t; \theta) \Delta(\tau_t; h) - \alpha \sum_{t=1}^n \gamma(v_t; \theta) \Delta(\tau_{t-1}; h) \\ &= -\frac{K(0)}{nh} \sum_{t=1}^n v_t \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t (1 + o_P(1)) = -\frac{K(0) \cdot (1 + o_P(1))}{nh} \sum_{t=1}^n v_t \gamma(v_t; \theta) w_t. \end{aligned}$$

We are now ready to complete the proofs of Theorems 1 and 2.

Proof of Theorem 1.

Observe that

$$\mathbb{E}_\star[\lambda | X_n, Y_n] = \frac{\iint \lambda e^{l_n(\lambda, \theta)} \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\iint e^{l_n(\lambda, \theta)} \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta} = \frac{\iint \lambda L_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\iint L_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta} = \frac{q_1(X_n, Y_n)}{p_1(X_n, Y_n)}, \quad (6)$$

where $q_1(X_n, Y_n) = \iint \lambda L_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta$ and $p_1(X_n, Y_n) = \iint L_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta$.

$$\mathbb{E}[\lambda] = \int \lambda \pi_\lambda(\lambda) d\lambda = \frac{\int \lambda \pi_\lambda(\lambda) d\lambda}{\int \pi_\lambda(\lambda) d\lambda} \frac{\int G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} = \frac{\iint \lambda G_n(\theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\iint G_n(\theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta} = \frac{q(X_n, Y_n)}{p(X_n, Y_n)}, \quad (7)$$

where $q(X_n, Y_n) = \iint \lambda G_n(\theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta$ and $p(X_n, Y_n) = \iint G_n(\theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta$.

Based on (6) and (7), we have that

$$\begin{aligned} q_1(X_n, Y_n) - q(X_n, Y_n) &= \iint \lambda (L_n(\lambda, \theta) - G_n(\theta)) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta, \\ p_1(X_n, Y_n) - p(X_n, Y_n) &= \iint (L_n(\lambda, \theta) - G_n(\theta)) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta. \end{aligned}$$

Therefore, by equation (2) we have

$$\begin{aligned} \mathbb{E}_\star[\lambda | X_n, Y_n] - \mathbb{E}[\lambda] &= \frac{q_1(X_n, Y_n)}{p_1(X_n, Y_n)} - \frac{q(X_n, Y_n)}{p(X_n, Y_n)} \\ &= \frac{1}{p_1(X_n, Y_n) p(X_n, Y_n)} [q_1(X_n, Y_n) p(X_n, Y_n) - p_1(X_n, Y_n) q(X_n, Y_n)] \end{aligned}$$

$$\begin{aligned}
&= \frac{1}{p_1(X_n, Y_n) p(X_n, Y_n)} [(q_1(X_n, Y_n) - q(X_n, Y_n)) p(X_n, Y_n) - (p_1(X_n, Y_n) - p(X_n, Y_n)) q(X_n, Y_n)] \\
&= \frac{1}{p_1(X_n, Y_n)} [(q_1(X_n, Y_n) - q(X_n, Y_n)) - (p_1(X_n, Y_n) - p(X_n, Y_n)) \mathbb{E}[\lambda]] \\
&= \frac{1}{p_1(X_n, Y_n)} \iint (\lambda - \mathbb{E}[\lambda]) (L_n(\lambda, \theta) - G_n(\theta)) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta \\
&= \frac{1 + o_P(1)}{p_1(X_n, Y_n)} \iint (\lambda - \mathbb{E}[\lambda]) G_n(\theta) \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta \\
&= (1 + o_P(1)) \int \frac{G_n(\theta)}{p_1(X_n, Y_n)} \left(\int (\lambda - \mathbb{E}[\lambda]) \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) \pi_\lambda(\lambda) d\lambda \right) \pi_\theta(\theta) d\theta,
\end{aligned}$$

where

$$\begin{aligned}
p_1(X_n, Y_n) &= \iint L_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta \\
&= \iint G_n(\theta) \left(\sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) + 1 \right) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta \\
&= \int G_n(\theta) \left(1 + \int \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) \pi_\lambda(\lambda) d\lambda \right) \pi_\theta(\theta) d\theta.
\end{aligned}$$

Therefore, it follows that

$$\begin{aligned}
\mathbb{E}_\star[\lambda | X_n, Y_n] - \mathbb{E}[\lambda] &= \frac{\int G_n(\theta) \left(\int (\lambda - \mathbb{E}[\lambda]) \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) \pi_\lambda(\lambda) d\lambda \right) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \left(1 + \int \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) \pi_\lambda(\lambda) d\lambda \right) \pi_\theta(\theta) d\theta} \\
&= \frac{\int G_n(\theta) R_{2n}(\lambda, \theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) R_{1n}(\lambda, \theta) \pi_\theta(\theta) d\theta},
\end{aligned}$$

where $R_{1n}(\lambda, \theta) = 1 + \int \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) \pi_\lambda(\lambda) d\lambda$ and

$$R_{2n}(\lambda, \theta) = \int (\lambda - \mathbb{E}[\lambda]) \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) \pi_\lambda(\lambda) d\lambda.$$

Equivalently, we have

$$\int G_n(\theta) R_{1n}(\lambda, \theta) \pi_\theta(\theta) d\theta \cdot (\mathbb{E}_\star[\lambda | X_n, Y_n] - \mathbb{E}[\lambda]) = \int G_n(\theta) R_{2n}(\lambda, \theta) \pi_\theta(\theta) d\theta. \quad (8)$$

Divide (8) by $\int G_n(\theta) \pi_\theta(\theta) d\theta$. Then we have

$$\frac{\int G_n(\theta) R_{1n}(\lambda, \theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \cdot (\mathbb{E}_\star[\lambda | X_n, Y_n] - \mathbb{E}[\lambda]) = \frac{\int G_n(\theta) R_{2n}(\lambda, \theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta}. \quad (9)$$

We now express $R_{2n}(\lambda, \theta)$ as follows:

$$R_{2n}(\lambda, \theta) = \int (\lambda - \mathbb{E}[\lambda]) \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) \pi_\lambda(\lambda) d\lambda$$

$$\begin{aligned}
&= (1 + o_P(1)) \int (\lambda - \mathbb{E}[\lambda]) \left(-\frac{K(0)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t v_t \right) \pi_\lambda(\lambda) d\lambda \\
&= -(1 + o_P(1)) K(0) \frac{1}{na_n} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t v_t \int (\lambda - \mathbb{E}[\lambda]) \frac{1}{\lambda} \pi_\lambda(\lambda) d\lambda \\
&= -(1 + o_P(1)) K(0) \frac{1}{na_n} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t v_t (1 - \mathbb{E}[\lambda] \mathbb{E}[\lambda^{-1}]) + o_P(1) \\
&= -(1 + o_P(1)) K(0) (1 - \mathbb{E}[\lambda] \mathbb{E}[\lambda^{-1}]) \frac{1}{na_n} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t v_t + o_P(1) \\
&= \frac{d_1}{na_n} \sum_{t=1}^n v_t \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t = \frac{d_1 (1 + o_P(1))}{na_n} \sum_{t=1}^n v_t \gamma(v_t; \theta) w_t,
\end{aligned}$$

where $d_1 = -K(0)(1 - \mathbb{E}[\lambda]\mathbb{E}[\lambda^{-1}])$.

By ignoring the high order $o_P(1)$ in the following derivations, we therefore have

$$\begin{aligned}
\int G_n(\theta) R_{2n}(\lambda, \theta) \pi_\theta(\theta) d\theta &= \int G_n(\theta) \left[\frac{d_1}{na_n} \sum_{t=1}^n \gamma(v_t; \theta) w_t v_t \right] \pi_\theta(\theta) d\theta \\
&= \frac{d_1}{na_n} \sum_{t=1}^n w_t v_t \int \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta.
\end{aligned} \tag{10}$$

Divide (10) by $\int G_n(\theta) \pi_\theta(\theta) d\theta$. It then follows that

$$\frac{\int G_n(\theta) R_{2n}(\lambda, \theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} = \frac{d_1}{na_n} \sum_{t=1}^n w_t v_t \frac{\int \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta}.$$

We know that

$$G_n(\theta) = \prod_{t=1}^n f(v_t; \theta) = e^{\sum_{t=1}^n \log f(v_t; \theta)} = e^{n \mathbb{E}_\theta [\log f(v_1; \theta)]} (1 + o_P(1)), \tag{11}$$

where $\mathbb{E}_\theta [\log f(v_1; \theta)] = \int \log f(v_1; \theta) f(v; \theta) dv$.

Let $A_n(\theta) = \frac{e^{n \mathbb{E}_\theta [\log f(v_1; \theta)]}}{\int e^{n \mathbb{E}_\theta [\log f(v_1; \theta)]} \pi_\theta(\theta) d\theta}$ and $\phi_n(v_t) = \int \gamma(v_t; \theta) A_n(\theta) \pi_\theta(\theta) d\theta$. Then

$$\frac{\int G_n(\theta) R_{2n}(\lambda, \theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} = \frac{d_1}{na_n} \sum_{t=1}^n w_t v_t \phi_n(v_t). \tag{12}$$

Meanwhile, we have

$$\begin{aligned}
\int G_n(\theta) R_{1n}(\lambda, \theta) \pi_\theta(\theta) d\theta &= \int G_n(\theta) \left(1 + \int \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) \pi_\lambda(\lambda) d\lambda \right) \pi_\theta(\theta) d\theta \\
&= \int G_n(\theta) \pi_\theta(\theta) d\theta + \int G_n(\theta) \left(\int \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) \pi_\lambda(\lambda) d\lambda \right) \pi_\theta(\theta) d\theta \\
&= \int G_n(\theta) \pi_\theta(\theta) d\theta + \int G_n(\theta) \left(\int \left(-\frac{K(0)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) x_t^\top \Sigma_x^{-1} x_t v_t \right) \pi_\lambda(\lambda) d\lambda \right) \pi_\theta(\theta) d\theta
\end{aligned}$$

$$= \int G_n(\theta) \pi_\theta(\theta) d\theta - K(0) \mathbb{E}[\lambda^{-1}] \frac{1}{na_n} \sum_{t=1}^n w_t v_t \int \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta. \quad (13)$$

Divide (13) by $\int G_n(\theta) \pi_\theta(\theta) d\theta$. Then we have

$$\begin{aligned} \frac{\int G_n(\theta) R_{1n}(\lambda, \theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} &= 1 - K(0) \mathbb{E}[\lambda^{-1}] \frac{1}{na_n} \sum_{t=1}^n w_t v_t \frac{\int \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \\ &= 1 - K(0) \mathbb{E}[\lambda^{-1}] \frac{1}{na_n} \sum_{t=1}^n w_t v_t \phi_n(v_t). \end{aligned} \quad (14)$$

Based on (12) and (14), we have

$$\left(1 - K(0) \mathbb{E}[\lambda^{-1}] \frac{1}{na_n} \sum_{t=1}^n w_t v_t \phi_n(v_t)\right) \cdot (\mathbb{E}_\star[\lambda | X_n, Y_n] - \mathbb{E}[\lambda]) = \frac{d_1}{na_n} \sum_{t=1}^n w_t v_t \phi_n(v_t). \quad (15)$$

Denote $Z_n = \frac{1}{\sqrt{n}} \sum_{t=1}^n w_t (v_t \phi_n(v_t) - \mathbb{E}[v_1 \phi_n(v_1)])$. Since w_t and v_t are independent, we have

$\mathbb{E}[Z_n] = 0$ and

$$\begin{aligned} \mathbb{E}[Z_n^2] &= \frac{1}{n} \sum_{t=1}^n \mathbb{E}[w_t^2] \mathbb{E}\left[\left(v_t \phi_n(v_t) - \mathbb{E}[v_1 \phi_n(v_1)]\right)^2\right] \\ &= \mathbb{E}[w_1^2] \mathbb{E}\left[\left(v_t \phi_n(v_t) - \mathbb{E}[v_1 \phi_n(v_1)]\right)^2\right] \equiv \delta_n^2. \end{aligned}$$

By a standard central limit theorem for the i.i.d. random variable case, we have

$$\delta_n^{-1} Z_n \rightarrow_D \mathcal{N}(0, 1). \quad (16)$$

We then have $\left(1 - K(0) \mathbb{E}[\lambda^{-1}] \frac{1}{na_n} \sum_{t=1}^n w_t v_t \phi_n(v_t)\right) (\mathbb{E}_\star[\lambda | X_n, Y_n] - \mathbb{E}[\lambda]) = \frac{d_1}{na_n} \sum_{t=1}^n w_t v_t \phi_n(v_t) = \frac{d_1}{\sqrt{na_n}} (Z_n + \sqrt{n} \mathbb{E}[w_1] \mathbb{E}[v_1 \phi_n(v_1)])$.

Using $\frac{1}{n} \sum_{t=1}^n (w_t v_t \phi_n(v_t) - \mathbb{E}[w_1] \mathbb{E}[v_1 \phi_n(v_1)]) \rightarrow 0$ as $n \rightarrow \infty$, and recalling the definitions of $b_{1n} = a_n^{-1} \mathbb{E}[w_1] \mathbb{E}[v_1 \phi_n(v_1)]$ and $b_{0n} = 1 - K(0) \mathbb{E}[\lambda^{-1}] b_{1n}$ listed in Theorem 1, we obtain as $n \rightarrow \infty$

$$\sqrt{na_n} \delta_n^{-1} (b_{0n} (\mathbb{E}_\star[\lambda | X_n, Y_n] - \mathbb{E}[\lambda]) - d_1 b_{1n}) = \frac{d_1}{\delta_n} Z_n \rightarrow_D \mathcal{N}(0, d_1^2) \quad (17)$$

by Assumption 4 and equation (16).

Therefore, under Assumptions 1–4, when $m \rightarrow \infty$ and $n \rightarrow \infty$, we have

$$\begin{aligned} &\sqrt{n} \delta_n^{-1} a_n (b_{0n} (\hat{\lambda}_{mn} - \mathbb{E}[\lambda]) - d_1 b_{1n}) \\ &= \sqrt{n} \delta_n^{-1} a_n \left(\frac{b_{0n}}{m} \sum_{j=1}^m (\lambda_{jn} - \mathbb{E}_\star[\lambda | X_n, Y_n]) + b_{0n} (\mathbb{E}_\star[\lambda | X_n, Y_n] - \mathbb{E}[\lambda]) - d_1 b_{1n} \right) \end{aligned}$$

$$\begin{aligned}
&= \frac{\sqrt{n}\delta_n^{-1}a_nb_{0n}}{\sqrt{m}} \frac{1}{\sqrt{m}} \sum_{j=1}^m (\lambda_{jn} - \mathbb{E}_\star[\lambda|X_n, Y_n]) + \sqrt{n}\delta_n^{-1}a_n(b_{0n}(\mathbb{E}_\star[\lambda|X_n, Y_n] - \mathbb{E}[\lambda]) - d_1b_{1n}) \\
&= o_P(1) + \sqrt{n}\delta_n^{-1}a_n(b_{0n}(\mathbb{E}_\star[\lambda|X_n, Y_n] - \mathbb{E}[\lambda]) - d_1b_{1n}) \xrightarrow{D} \mathcal{N}(0, d_1^2).
\end{aligned} \tag{18}$$

Therefore, we have proved Theorem 1. In order to prove Theorem 2, we introduce a useful lemma in Lemma C below.

Lemma C. Under Assumption 5 listed in Appendix A1, we have as $n \rightarrow \infty$,

$$\Delta_n^{1/2}(\hat{\theta}_n)(\mathbb{E}[\theta|X_n, Y_n] - \mathbb{E}[\theta]) \xrightarrow{D} \mathcal{N}(0, \Sigma_0),$$

where Σ_0 is same as that in Assumption 5 (iv) and it can be estimated by a traditional method, such as the maximum likelihood estimation method.

Proof of Lemma C. By definition, we have that

$$\mathbb{E}[\theta|X_n, Y_n] = \int \theta f_n(\theta|X_n, Y_n) d\theta = \frac{\int \theta e^{g_n(\theta)} \pi_\theta(\theta) d\theta}{\int e^{g_n(\theta)} \pi_\theta(\theta) d\theta} = \frac{q_n(X_n, Y_n)}{p_n(X_n, Y_n)},$$

where $q_n(X_n, Y_n) = \int \theta e^{g_n(\theta)} \pi_\theta(\theta) d\theta$ and $p_n(X_n, Y_n) = \int e^{g_n(\theta)} \pi_\theta(\theta) d\theta$.

Taking the first order Taylor expansion of $g_n(\theta)$ in the neighborhood of $\hat{\theta}_n$, we have that

$$g_n(\theta) = g_n(\hat{\theta}_n) + g_n^{(1)}(\hat{\theta}_n)(\theta - \hat{\theta}_n) + \frac{1}{2}g_n^{(2)}(\hat{\theta}_n)(\theta - \hat{\theta}_n)^2 = g_n(\hat{\theta}_n) - \frac{1}{2}\Delta_n(\hat{\theta}_n)(\theta - \hat{\theta}_n)^2,$$

where $g_n^{(1)}(\hat{\theta}_n)$ and $g_n^{(2)}(\hat{\theta}_n)$ are the first order and second order derivatives of $g_n(\theta)$ evaluated at the point $\hat{\theta}_n$, respectively.

So $q_n(X_n, Y_n)$ can be written as

$$q_n(X_n, Y_n) = \int_{C_n} \theta e^{g_n(\hat{\theta}_n)} e^{-\frac{1}{2}\Delta_n(\hat{\theta}_n)(\theta - \hat{\theta}_n)^2} \pi_\theta(\theta) d\theta + \int_{D_n} \theta e^{g_n(\theta)} \pi_\theta(\theta) d\theta = q_{1n}(X_n, Y_n) + q_{2n}(X_n, Y_n),$$

where $q_{1n}(X_n, Y_n) = \int_{C_n} \theta e^{g_n(\hat{\theta}_n)} e^{-\frac{1}{2}\Delta_n(\hat{\theta}_n)(\theta - \hat{\theta}_n)^2} \pi_\theta(\theta) d\theta$ and $q_{2n}(X_n, Y_n) = \int_{D_n} \theta e^{g_n(\theta)} \pi_\theta(\theta) d\theta$.

Similarly, $p_n(X_n, Y_n)$ can be written as

$$p_n(X_n, Y_n) = \int_{C_n} e^{g_n(\hat{\theta}_n)} e^{-\frac{1}{2}\Delta_n(\hat{\theta}_n)(\theta - \hat{\theta}_n)^2} \pi_\theta(\theta) d\theta + \int_{D_n} e^{g_n(\theta)} \pi_\theta(\theta) d\theta = p_{1n}(X_n, Y_n) + p_{2n}(X_n, Y_n),$$

where $p_{1n}(X_n, Y_n) = \int_{C_n} e^{g_n(\hat{\theta}_n)} e^{-\frac{1}{2}\Delta_n(\hat{\theta}_n)(\theta - \hat{\theta}_n)^2} \pi_\theta(\theta) d\theta$ and $p_{2n}(X_n, Y_n) = \int_{D_n} e^{g_n(\theta)} \pi_\theta(\theta) d\theta$.

We first express $q_{1n}(X_n, Y_n)$ as follows:

$$\begin{aligned}
q_{1n}(X_n, Y_n) &= \int_{C_n} \theta e^{g_n(\hat{\theta}_n)} e^{-\frac{1}{2}\Delta_n(\hat{\theta}_n)(\theta - \hat{\theta}_n)^2} \pi_\theta(\theta) d\theta \\
&= \int_{|y| \leq c} (\hat{\theta}_n + \Delta_n^{-1/2}(\hat{\theta}_n)y) \Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \pi(\hat{\theta}_n + \Delta_n^{-1/2}(\hat{\theta}_n)y) e^{-\frac{1}{2}y^2} dy \\
&= \hat{\theta}_n \Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \int_{|y| \leq c} \pi(\hat{\theta}_n + \Delta_n^{-1/2}(\hat{\theta}_n)y) e^{-\frac{1}{2}y^2} dy \\
&\quad + \Delta_n^{-1/2}(\hat{\theta}_n) \Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \int_{|y| \leq c} \pi(\hat{\theta}_n + \Delta_n^{-1/2}(\hat{\theta}_n)y) y e^{-\frac{1}{2}y^2} dy \\
&= \hat{\theta}_n \Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \pi(\hat{\theta}_n) \int_{|y| \leq c} e^{-\frac{1}{2}y^2} dy + \hat{\theta}_n \Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \Delta_n^{-1/2}(\hat{\theta}_n) \pi^{(1)}(\hat{\theta}_n) \int_{|y| \leq c} y e^{-\frac{1}{2}y^2} dy \\
&\quad + \Delta_n^{-1/2}(\hat{\theta}_n) \Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \pi(\hat{\theta}_n) \int_{|y| \leq c} y e^{-\frac{1}{2}y^2} dy + o_P(\Delta_n^{-1}(\hat{\theta}_n)) \\
&= \hat{\theta}_n \Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \pi(\hat{\theta}_n) \int_{|y| \leq c} e^{-\frac{1}{2}y^2} dy + o_P(\Delta_n^{-1}(\hat{\theta}_n)). \tag{19}
\end{aligned}$$

The last equality is true, because

$$\begin{aligned}
\int_{|y| \leq c} y e^{-\frac{1}{2}y^2} dy &= \int_0^c y e^{-\frac{1}{2}y^2} dy + \int_{-c}^0 y e^{-\frac{1}{2}y^2} dy \\
&= \int_0^c y e^{-\frac{1}{2}y^2} dy + \int_c^0 (-y) e^{-\frac{1}{2}y^2} d(-y) = \int_0^c y e^{-\frac{1}{2}y^2} dy - \int_0^c y e^{-\frac{1}{2}y^2} dy = 0.
\end{aligned}$$

Similarly, we have

$$\begin{aligned}
p_{1n}(X_n, Y_n) &= \int_{C_n} e^{g_n(\hat{\theta}_n)} e^{-\frac{1}{2}\Delta_n(\hat{\theta}_n)(\theta - \hat{\theta}_n)^2} \pi_\theta(\theta) d\theta = \Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \int_{|y| \leq c} \pi(\hat{\theta}_n + \Delta_n^{-1/2}(\hat{\theta}_n)y) e^{-\frac{1}{2}y^2} dy \\
&= \Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \pi(\hat{\theta}_n) \int_{|y| \leq c} e^{-\frac{1}{2}y^2} dy + \Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \Delta_n^{-1/2}(\hat{\theta}_n) \pi^{(1)}(\hat{\theta}_n) \int_{|y| \leq c} y e^{-\frac{1}{2}y^2} dy \\
&\quad + o_P(\Delta_n^{-1}(\hat{\theta}_n)) = \Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \pi(\hat{\theta}_n) \int_{|y| \leq c} e^{-\frac{1}{2}y^2} dy + o_P(\Delta_n^{-1}(\hat{\theta}_n)). \tag{20}
\end{aligned}$$

Then, based on (19) and (20), we obtain

$$\frac{q_{1n}(X_n, Y_n)}{p_{1n}(X_n, Y_n)} = \hat{\theta}_n + \frac{o_P(\Delta_n^{-1}(\hat{\theta}_n))}{\Delta_n^{-1/2}(\hat{\theta}_n) e^{g_n(\hat{\theta}_n)} \pi(\hat{\theta}_n) \int_{|y| \leq c} e^{-\frac{1}{2}y^2} dy + o_P(\Delta_n^{-1}(\hat{\theta}_n))} = \hat{\theta}_n + o_P(\Delta_n^{-1/2}(\hat{\theta}_n)). \tag{21}$$

Note that under Assumption 5 (iii), we have

$$\frac{q_{2n}(X_n, Y_n)}{p_n(X_n, Y_n)} = o_P(\Delta_n^{-1/2}(\hat{\theta}_n)) \text{ and } \frac{p_{2n}(X_n, Y_n)}{p_n(X_n, Y_n)} = o_P(\Delta_n^{-1/2}(\hat{\theta}_n)). \tag{22}$$

Therefore, in view of (21) and (22), with Assumption 5 (iii), it is shown that

$$|\mathbb{E}[\theta | X_n, Y_n] - \hat{\theta}_n| = \left| \frac{q_n(X_n, Y_n)}{p_n(X_n, Y_n)} - \hat{\theta}_n \right| = \left| \frac{q_{1n}(X_n, Y_n)}{p_n(X_n, Y_n)} + \frac{q_{2n}(X_n, Y_n)}{p_n(X_n, Y_n)} - \hat{\theta}_n \right|$$

$$\begin{aligned}
&\leq \left| \frac{q_{1n}(X_n, Y_n)}{p_n(X_n, Y_n)} - \hat{\theta}_n \right| + \left| \frac{q_{2n}(X_n, Y_n)}{p_n(X_n, Y_n)} \right| \\
&= \left| \frac{q_{1n}(X_n, Y_n)}{p_n(X_n, Y_n)} - \frac{q_{1n}(X_n, Y_n)}{p_{1n}(X_n, Y_n)} + \frac{q_{1n}(X_n, Y_n)}{p_{1n}(X_n, Y_n)} - \hat{\theta}_n \right| + \left| \frac{q_{2n}(X_n, Y_n)}{p_n(X_n, Y_n)} \right| \\
&\leq \left| \frac{q_{1n}(X_n, Y_n)}{p_n(X_n, Y_n)} - \frac{q_{1n}(X_n, Y_n)}{p_{1n}(X_n, Y_n)} \right| + \left| \frac{q_{1n}(X_n, Y_n)}{p_{1n}(X_n, Y_n)} - \hat{\theta}_n \right| + \left| \frac{q_{2n}(X_n, Y_n)}{p_n(X_n, Y_n)} \right| \\
&= \left| q_{1n}(X_n, Y_n) \right| \left| \frac{1}{p_n(X_n, Y_n)} - \frac{1}{p_{1n}(X_n, Y_n)} \right| + \left| \frac{q_{1n}(X_n, Y_n)}{p_{1n}(X_n, Y_n)} - \hat{\theta}_n \right| + \left| \frac{q_{2n}(X_n, Y_n)}{p_n(X_n, Y_n)} \right| \\
&= \left| q_{1n}(X_n, Y_n) \right| \left| \frac{p_{2n}(X_n, Y_n)}{p_n(X_n, Y_n)p_{1n}(X_n, Y_n)} \right| + o_P(\Delta_n^{-1/2}(\hat{\theta}_n)) \\
&= \left| \frac{q_{1n}(X_n, Y_n)}{p_{1n}(X_n, Y_n)} \right| \left| \frac{p_{2n}(X_n, Y_n)}{p_n(X_n, Y_n)} \right| + o_P(\Delta_n^{-1/2}(\hat{\theta}_n)) = o_P(\Delta_n^{-1/2}(\hat{\theta}_n)).
\end{aligned}$$

Thus, $\Delta_n^{1/2}(\hat{\theta}_n)(\mathbb{E}[\theta|X_n, Y_n] - \hat{\theta}_n) \rightarrow_P 0$.

Therefore, with Assumption 5(iv), we obtain

$$\Delta_n^{1/2}(\hat{\theta}_n)(\mathbb{E}[\theta|X_n, Y_n] - \mathbb{E}[\theta]) = \Delta_n^{1/2}(\hat{\theta}_n)(\mathbb{E}[\theta|X_n, Y_n] - \hat{\theta}_n) + \Delta_n^{1/2}(\hat{\theta}_n)(\hat{\theta}_n - \mathbb{E}[\theta]) \rightarrow_D \mathcal{N}(0, \Sigma_0), \quad (23)$$

which completes the proof of Lemma C.

We next give the proof of Theorem 2 by Lemma C.

Proof of Theorem 2(i).

From equation (1) in the proof of Theorem 1, we have obtained that

$$l_n(\lambda, \theta) = \log L_n(\lambda, \theta) = \sum_{t=1}^n \log f(\hat{v}_t; \theta) = \sum_{t=1}^n \log f(v_t; \theta) + \sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) (1 + o_P(1)),$$

where $\gamma(v_t; \theta) = \frac{f^{(1)}(v_t; \theta)}{f(v_t; \theta)}$.

Therefore, we have

$$e^{l_n(\lambda, \theta)} = e^{\sum_{t=1}^n \log f(v_t; \theta)} e^{\sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) (1 + o_P(1))} = e^{\sum_{t=1}^n \log f(v_t; \theta)} + R_n(\lambda, \theta),$$

where $R_n(\lambda, \theta) = (e^{\sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) + o_P(1)} - 1) e^{\sum_{t=1}^n \log f(v_t; \theta)} = (\sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h)) G_n(\theta)$.

By ignoring the high order $o_P(1)$ in the following derivations, we then have the following two equations:

$$\begin{aligned}
\iint e^{l_n(\lambda, \theta)} \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta &= \iint e^{\sum_{t=1}^n \log f(v_t; \theta)} \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta + \iint R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta \\
&= \int G_n(\theta) \pi_\theta(\theta) d\theta + \iint R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta,
\end{aligned} \quad (24)$$

$$\begin{aligned}
\iint \theta e^{l_n(\lambda, \theta)} \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta &= \iint \theta e^{\sum_{t=1}^n \log f(v_t; \theta)} \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta + \iint \theta R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta \\
&= \int \theta G_n(\theta) \pi_\theta(\theta) d\theta + \iint \theta R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta.
\end{aligned} \tag{25}$$

It follows from (24) and (25) that

$$\begin{aligned}
\mathbb{E}_\star[\theta | X_n, Y_n] &= \frac{\iint \theta e^{l_n(\lambda, \theta)} \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\iint e^{l_n(\lambda, \theta)} \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta} \\
&= \frac{\int \theta G_n(\theta) \pi_\theta(\theta) d\theta + \iint \theta R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta + \iint R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}.
\end{aligned}$$

As $\mathbb{E}[\theta | X_n, Y_n] = \frac{\int \theta G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta}$, we have

$$\begin{aligned}
&\mathbb{E}_\star[\theta | X_n, Y_n] - \mathbb{E}[\theta | X_n, Y_n] \\
&= \frac{\int \theta G_n(\theta) \pi_\theta(\theta) d\theta + \iint \theta R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta + \iint R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta} - \frac{\int \theta G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \\
&= \frac{\iint \theta R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} - \mathbb{E}[\theta | X_n, Y_n] \frac{\iint R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta}.
\end{aligned} \tag{26}$$

From the proof of Theorem 1, we have obtained that

$$\sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h) = -\frac{K(0)}{nh} \sum_{t=1}^n v_t \gamma(v_t; \theta) w_t (1 + o_P(1)).$$

Recall that $R_n(\lambda, \theta) = (\sum_{t=1}^n \gamma(v_t; \theta) \Gamma(\tau_t; \alpha, h)) G_n(\theta)$. Therefore, we obtain

$$\begin{aligned}
&\frac{\iint R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} = -\frac{\iint \left(\frac{K(0)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) w_t v_t\right) G_n(\theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \\
&= -K(0) \mathbb{E}[\lambda^{-1}] \frac{\int \left(\frac{1}{na_n} \sum_{t=1}^n \gamma(v_t; \theta) w_t v_t\right) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \\
&= -K(0) \mathbb{E}[\lambda^{-1}] \frac{1}{na_n} \sum_{t=1}^n w_t v_t \frac{\int \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta}.
\end{aligned}$$

Similarly, we can have

$$\begin{aligned}
&\frac{\iint \theta R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} = -\frac{\iint \theta \left(\frac{K(0)}{nh} \sum_{t=1}^n \gamma(v_t; \theta) w_t v_t\right) G_n(\theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \\
&= -K(0) \mathbb{E}[\lambda^{-1}] \frac{\int \left(\frac{1}{na_n} \sum_{t=1}^n \gamma(v_t; \theta) w_t v_t\right) \theta G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \\
&= -K(0) \mathbb{E}[\lambda^{-1}] \frac{1}{na_n} \sum_{t=1}^n w_t v_t \frac{\int \theta \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta}.
\end{aligned}$$

Recall that

$$G_n(\theta) = \prod_{t=1}^n f(v_t; \theta) = e^{\sum_{t=1}^n \log f(v_t; \theta)} = e^{n\mathbb{E}_\theta[\log f(v_1; \theta)]}(1 + o_P(1)), \quad (27)$$

where $\mathbb{E}_\theta[\log f(v_1; \theta)] = \int \log f(v_1; \theta) f(v; \theta) d\nu$. Denote $A_n(\theta) = \frac{e^{n\mathbb{E}_\theta[\log f(v_1; \theta)]}}{\int e^{n\mathbb{E}_\theta[\log f(v_1; \theta)]} \pi_\theta(\theta) d\theta}$.

Simple decompositions give

$$\begin{aligned} & \mathbb{E}_\star[\theta | X_n, Y_n] - \mathbb{E}[\theta | X_n, Y_n] \\ &= \frac{\iint \theta R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} - \mathbb{E}[\theta | X_n, Y_n] \frac{\iint R_n(\lambda, \theta) \pi_\lambda(\lambda) \pi_\theta(\theta) d\lambda d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \\ &= -K(0)\mathbb{E}[\lambda^{-1}] \frac{1}{a_n n} \sum_{t=1}^n w_t v_t \frac{\int \theta \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \\ &\quad + \mathbb{E}[\theta | X_n, Y_n] K(0)\mathbb{E}[\lambda^{-1}] \frac{1}{a_n n} \sum_{t=1}^n w_t v_t \frac{\int \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \\ &= -K(0)\mathbb{E}[\lambda^{-1}] \frac{1}{a_n n} \sum_{t=1}^n w_t v_t \left(\frac{\int \theta \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} - \mathbb{E}[\theta | X_n, Y_n] \frac{\int \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \right) \\ &= -K(0)\mathbb{E}[\lambda^{-1}] \frac{1}{a_n n} \sum_{t=1}^n w_t v_t \frac{\int (\theta - \mathbb{E}[\theta | X_n, Y_n]) \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \\ &= K(0)\mathbb{E}[\lambda^{-1}] \frac{1}{a_n n} \sum_{t=1}^n w_t v_t \left(\int (\mathbb{E}[\theta] - \theta) \gamma(v_t; \theta) A_n(\theta) \pi_\theta(\theta) d\theta \right) \\ &\quad + K(0)\mathbb{E}[\lambda^{-1}] \frac{1}{a_n n} \sum_{t=1}^n w_t v_t \left(\int \gamma(v_t; \theta) A_n(\theta) \pi_\theta(\theta) d\theta \right) \cdot (\mathbb{E}[\theta | X_n, Y_n] - \mathbb{E}[\theta]). \end{aligned} \quad (28)$$

By Lemma C, we have $\mathbb{E}[\theta | X_n, Y_n] - \mathbb{E}[\theta] = o_P(\Delta_n^{-1/2}(\hat{\theta}_n))$. Therefore, it follows that

$$\begin{aligned} & \mathbb{E}_\star[\theta | X_n, Y_n] - \mathbb{E}[\theta | X_n, Y_n] = K(0)\mathbb{E}[\lambda^{-1}] (1 + o_P(1)) \\ &\quad \times \frac{1}{na_n} \sum_{t=1}^n w_t v_t \left(\int (\mathbb{E}[\theta] - \theta) \gamma(v_t; \theta) A_n(\theta) \pi_\theta(\theta) d\theta \right). \end{aligned}$$

Let $\gamma_n(v_t) = \int (\mathbb{E}[\theta] - \theta) \gamma(v_t; \theta) A_n(\theta) \pi_\theta(\theta) d\theta$. Denote $S_n = \frac{1}{\sqrt{n}} \sum_{t=1}^n w_t (v_t \gamma_n(v_t) - \mathbb{E}[v_1 \gamma_n(v_1)])$.

Since w_t and v_t are independent, we have $\mathbb{E}[S_n] = 0$ and

$$\begin{aligned} \mathbb{E}[S_n S_n^\top] &= \frac{1}{n} \sum_{t=1}^n \mathbb{E}[w_t^2] \mathbb{E}\left[(v_t \gamma_n(v_t) - \mathbb{E}[v_1 \gamma_n(v_1)]) (v_t \gamma_n(v_t) - \mathbb{E}[v_1 \gamma_n(v_1)])^\top \right] \\ &= \mathbb{E}[w_t^2] \mathbb{E}\left[(v_t \gamma_n(v_t) - \mathbb{E}[v_1 \gamma_n(v_1)]) (v_t \gamma_n(v_t) - \mathbb{E}[v_1 \gamma_n(v_1)])^\top \right] \equiv \sigma_n^2. \end{aligned}$$

A standard central limit theorem implies that as $n \rightarrow \infty$

$$\sigma_n^{-1} S_n \rightarrow_D \mathcal{N}(0, I_p), \quad (29)$$

where I_p denotes the $p \times p$ identity matrix and p denotes the number of parameters in the vector θ .

Let $T_n = \mathbb{E}_\star[\theta|X_n, Y_n] - \mathbb{E}[\theta|X_n, Y_n] = K(0)\mathbb{E}[\lambda^{-1}] \frac{1}{\sqrt{n}a_n} (S_n + \sqrt{n}\mathbb{E}[w_1]\mathbb{E}[\nu_1\gamma_n(\nu_1)])$. Based on (29), under Assumption 3 (iv), we obtain

$$\sqrt{n}\sigma_n^{-1}a_n \left(\frac{1}{K(0)\mathbb{E}[\lambda^{-1}]} T_n - a_n^{-1}\mathbb{E}[w_1]\mathbb{E}[\nu_1\gamma_n(\nu_1)] \right) = \sigma_n^{-1}S_n \rightarrow_D \mathcal{N}(0, I_p). \quad (30)$$

Equivalently, we have

$$\begin{aligned} & \sqrt{n}\sigma_n^{-1}a_n (\mathbb{E}_\star[\theta|X_n, Y_n] - \mathbb{E}[\theta|X_n, Y_n] - a_n^{-1}K(0)\mathbb{E}[\lambda^{-1}]\mathbb{E}[w_1]\mathbb{E}[\nu_1\gamma_n(\nu_1)]) \\ & \rightarrow_D \mathcal{N}(0, K^2(0)\mathbb{E}^2[\lambda^{-1}]I_p). \end{aligned} \quad (31)$$

By a conditional central limit theorem for the i.i.d. case, we have as $m \rightarrow \infty$

$$\frac{1}{\sqrt{m}} \sum_{j=1}^m (\theta_{jn} - \mathbb{E}_\star[\theta|X_n, Y_n]) = O_P(1).$$

Let $b_{2n} = a_n^{-1}K(0)\mathbb{E}[\lambda^{-1}]\mathbb{E}[w_1]\mathbb{E}[\nu_1\gamma_n(\nu_1)]$.

Then, under Assumptions 1–3 and 5, we obtain as $m \rightarrow \infty$ and $n \rightarrow \infty$

$$\begin{aligned} & \sqrt{n}\sigma_n^{-1}a_n (\hat{\theta}_{mn} - \mathbb{E}[\theta] - b_{2n}) \\ &= \sqrt{n}\sigma_n^{-1}a_n \left(\frac{1}{m} \sum_{j=1}^m \theta_{jn} - \mathbb{E}_\star[\theta|X_n, Y_n] + \mathbb{E}_\star[\theta|X_n, Y_n] - \mathbb{E}[\theta|X_n, Y_n] - b_{2n} + \mathbb{E}[\theta|X_n, Y_n] - \mathbb{E}[\theta] \right) \\ &= \sqrt{n}\sigma_n^{-1}a_n \left(\frac{1}{m} \sum_{j=1}^m \theta_{jn} - \mathbb{E}_\star[\theta|X_n, Y_n] \right) + \sqrt{n}\sigma_n^{-1}a_n (\mathbb{E}_\star[\theta|X_n, Y_n] - \mathbb{E}[\theta|X_n, Y_n] - b_{2n}) \\ &\quad + \sqrt{n}\sigma_n^{-1}a_n \Delta_n^{-1/2}(\hat{\theta}_n) \Delta_n^{1/2}(\hat{\theta}_n) (\mathbb{E}[\theta|X_n, Y_n] - \mathbb{E}[\theta]) \\ &= \frac{\sqrt{n}\sigma_n^{-1}a_n}{\sqrt{m}} \frac{1}{\sqrt{m}} \sum_{j=1}^m (\theta_{jn} - \mathbb{E}_\star[\theta|X_n, Y_n]) + \sqrt{n}\sigma_n^{-1}a_n (\mathbb{E}_\star[\theta|X_n, Y_n] - \mathbb{E}[\theta|X_n, Y_n] - b_{2n}) + o_P(1) \\ &\rightarrow_D \mathcal{N}(0, K^2(0)\mathbb{E}^2[\lambda^{-1}]I_p). \end{aligned}$$

To sum up, under Assumptions 1–3 and 5, as $m \rightarrow \infty$ and $n \rightarrow \infty$, we have

$$\sqrt{n}\sigma_n^{-1}a_n (\hat{\theta}_{mn} - \mathbb{E}[\theta] - b_{2n}) \rightarrow_D \mathcal{N}(0, K^2(0)\mathbb{E}^2[\lambda^{-1}]I_p),$$

which completes the proof of Theorem 2(i).

Proof of Theorem 2(ii).

Equation (28) implies

$$\begin{aligned}
& \mathbb{E}_\star[\theta|X_n, Y_n] - \mathbb{E}[\theta|X_n, Y_n] = -K(0)\mathbb{E}[\lambda^{-1}] \frac{1}{a_n n} \sum_{t=1}^n w_t v_t \frac{\int (\theta - \mathbb{E}[\theta|X_n, Y_n]) \gamma(v_t; \theta) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \\
&= -\frac{K(0)\mathbb{E}[\lambda^{-1}]}{a_n n} \sum_{t=1}^n v_t \gamma(v_t) w_t \left(\frac{\int (\theta - \mathbb{E}[\theta|X_n, Y_n]) G_n(\theta) \pi_\theta(\theta) d\theta}{\int G_n(\theta) \pi_\theta(\theta) d\theta} \right) \\
&= -\frac{K(0)\mathbb{E}[\lambda^{-1}]}{a_n n} \sum_{t=1}^n v_t \gamma(v_t) w_t \cdot (\mathbb{E}[\theta|X_n, Y_n] - \mathbb{E}[\theta|X_n, Y_n]) = 0
\end{aligned}$$

when $\gamma(v; \theta) \equiv \gamma(v)$.

Therefore, we have as $m \rightarrow \infty$ and $n \rightarrow \infty$

$$\begin{aligned}
& \Delta_n^{1/2}(\hat{\theta}_n) (\hat{\theta}_{mn} - \mathbb{E}[\theta]) \\
&= \Delta_n^{1/2}(\hat{\theta}_n) \left(\frac{1}{m} \sum_{j=1}^m \theta_{jn} - \mathbb{E}_\star[\theta|X_n, Y_n] + \mathbb{E}_\star[\theta|X_n, Y_n] - \mathbb{E}[\theta|X_n, Y_n] + \mathbb{E}[\theta|X_n, Y_n] - \mathbb{E}[\theta] \right) \\
&= \Delta_n^{1/2}(\hat{\theta}_n) \left(\frac{1}{m} \sum_{j=1}^m \theta_{jn} - \mathbb{E}_\star[\theta|X_n, Y_n] \right) + \Sigma_n^{1/2}(\hat{\theta}_n) (\mathbb{E}_\star[\theta|X_n, Y_n] - \mathbb{E}[\theta|X_n, Y_n]) \\
&\quad + \Delta_n^{1/2}(\hat{\theta}_n) (\mathbb{E}[\theta|X_n, Y_n] - \mathbb{E}[\theta]) \\
&= \frac{\Delta_n^{1/2}(\hat{\theta}_n)}{\sqrt{m}} \frac{1}{\sqrt{m}} \sum_{j=1}^m (\theta_{jn} - \mathbb{E}_\star[\theta|X_n, Y_n]) + \Delta_n^{1/2}(\hat{\theta}_n) (\mathbb{E}[\theta|X_n, Y_n] - \mathbb{E}[\theta]) \\
&= \Delta_n^{1/2}(\hat{\theta}_n) (\mathbb{E}[\theta|X_n, Y_n] - \mathbb{E}[\theta]) + o_P(1) \rightarrow_D \mathcal{N}(0, \Sigma_0),
\end{aligned}$$

which completes the proof of Theorem 2(ii).

Appendix D: Plots for $\hat{\beta}_1(\tau)$ and $\hat{\beta}_2(\tau)$ in simulation

Figure 1: Median of $\hat{\beta}_1(\tau)$ and $\hat{\beta}_2(\tau)$ based on 1000 replications under DGP1 in Case 1 with sample size $n = 200, 600$ and 1200 (Graphs of $\hat{\beta}_1(\tau)$ are displayed in the first row and graphs of $\hat{\beta}_2(\tau)$ are displayed in the second row).

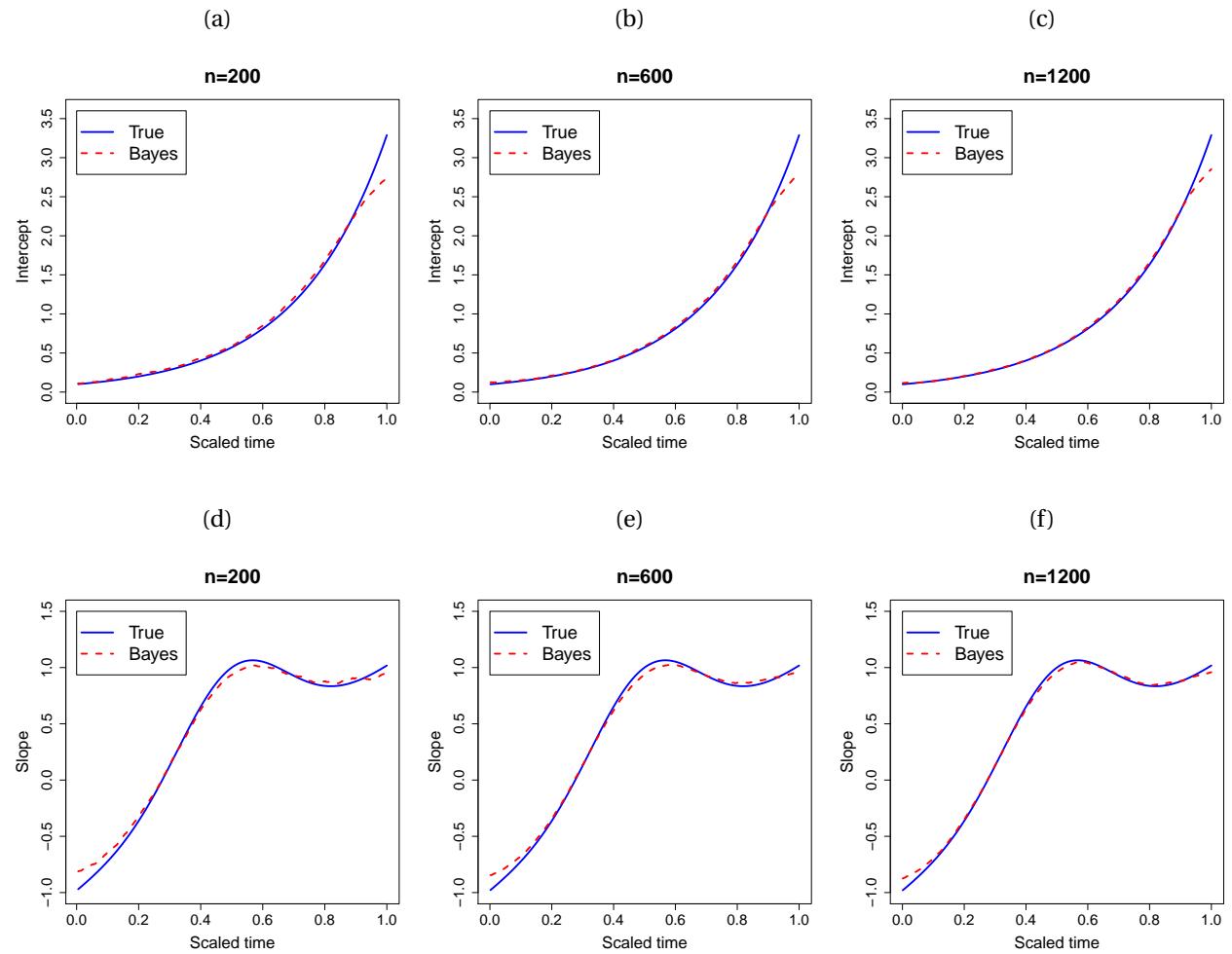


Figure 2: Median of $\hat{\beta}_1(\tau)$ and $\hat{\beta}_2(\tau)$ based on 1000 replications under DGP2 in Case 1 with sample size $n = 200, 600$ and 1200 (Graphs of $\hat{\beta}_1(\tau)$ are displayed in the first row and graphs of $\hat{\beta}_2(\tau)$ are displayed in the second row).

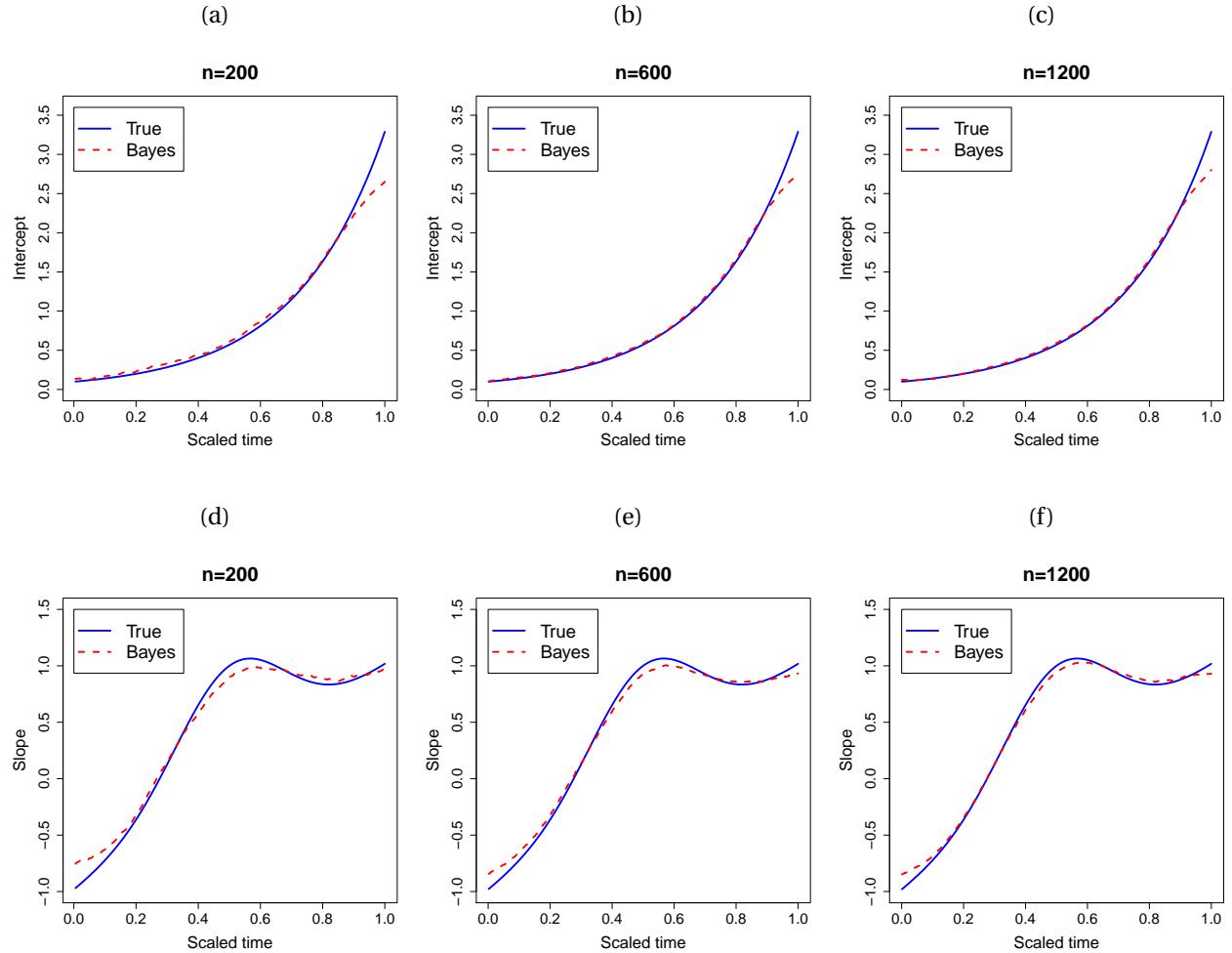


Figure 3: Median of $\hat{\beta}_1(\tau)$ and $\hat{\beta}_2(\tau)$ based on 1000 replications under DGP3 in Case 1 with sample size $n = 200, 600$ and 1200 (Graphs of $\hat{\beta}_1(\tau)$ are displayed in the first row and graphs of $\hat{\beta}_2(\tau)$ are displayed in the second row).

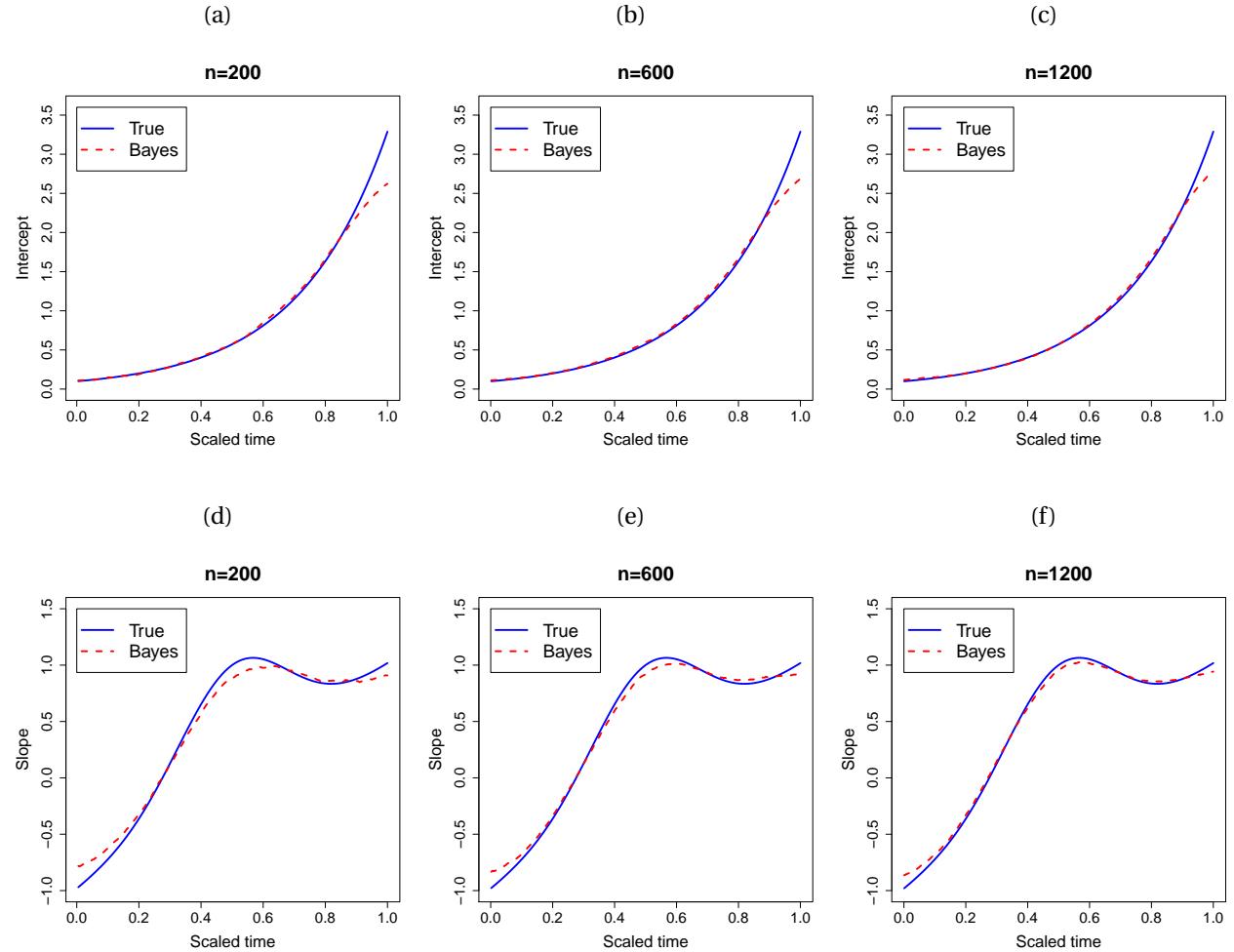


Figure 4: Median of $\hat{\beta}_1(\tau)$ and $\hat{\beta}_2(\tau)$ based on 1000 replications under DGP1 in Case 2 with sample size $n = 200, 600$ and 1200 (Graphs of $\hat{\beta}_1(\tau)$ are displayed in the first row and graphs of $\hat{\beta}_2(\tau)$ are displayed in the second row).

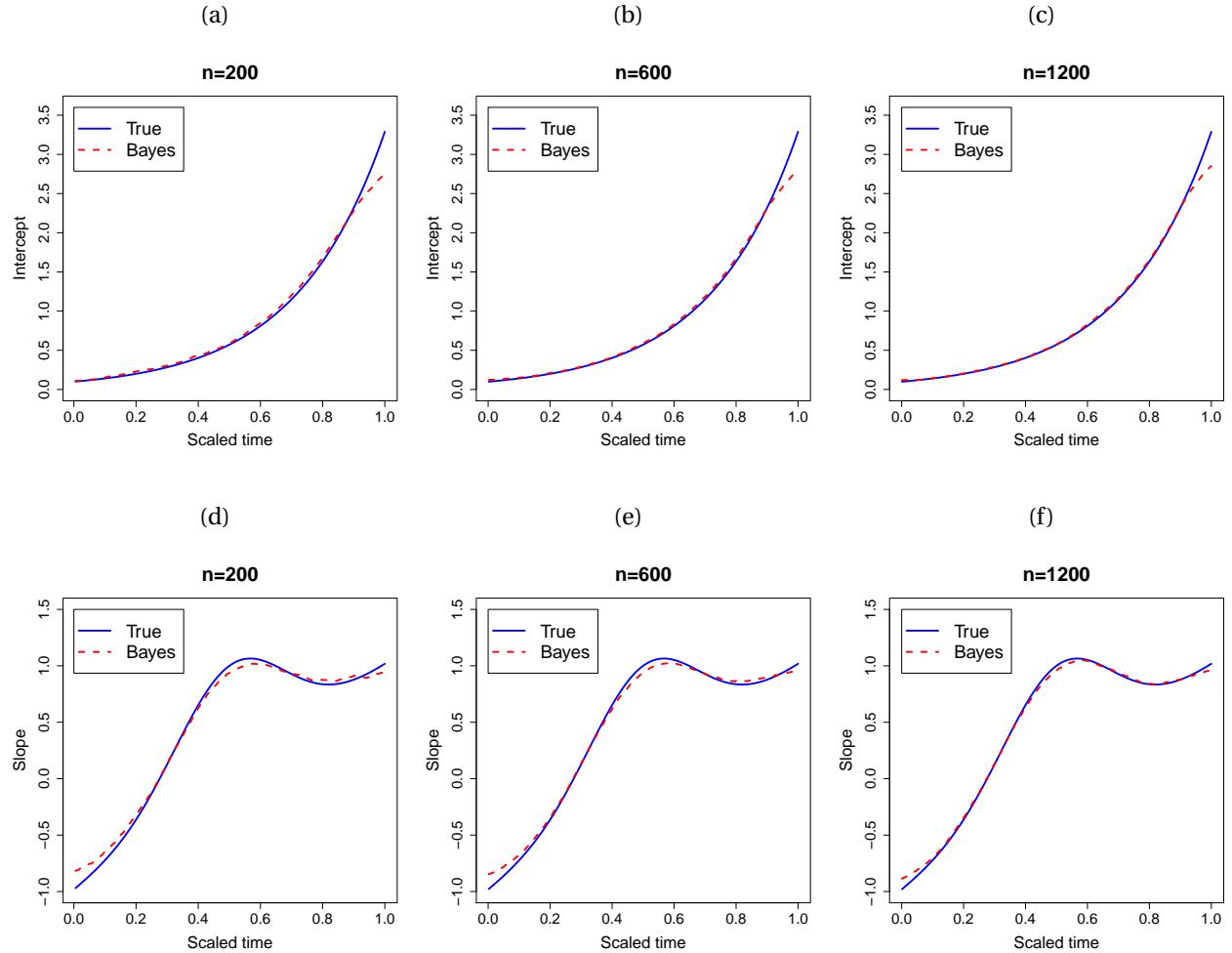


Figure 5: Median of $\hat{\beta}_1(\tau)$ and $\hat{\beta}_2(\tau)$ based on 1000 replications under DGP2 in Case 2 with sample size $n = 200, 600$ and 1200 (Graphs of $\hat{\beta}_1(\tau)$ are displayed in the first row and graphs of $\hat{\beta}_2(\tau)$ are displayed in the second row).

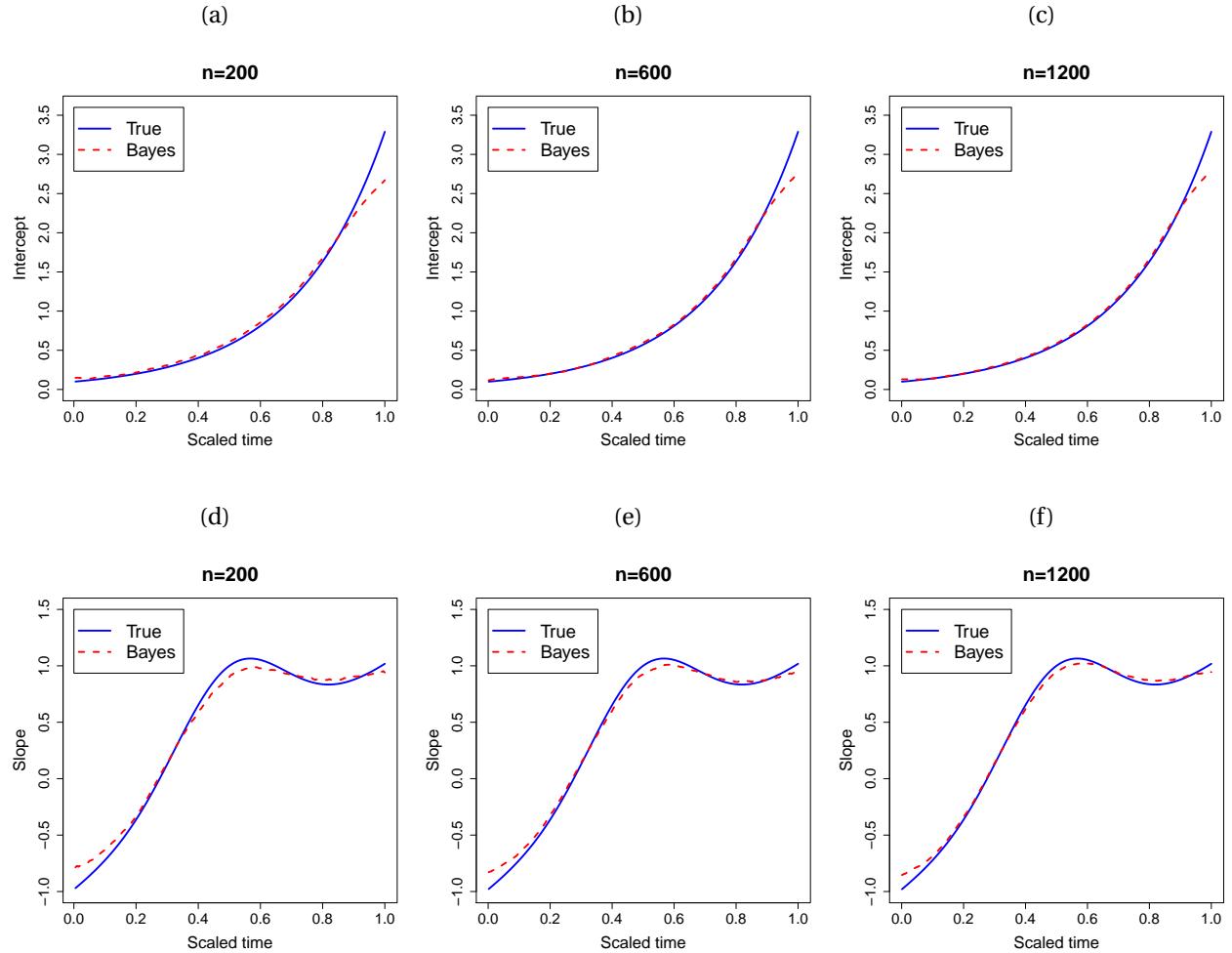
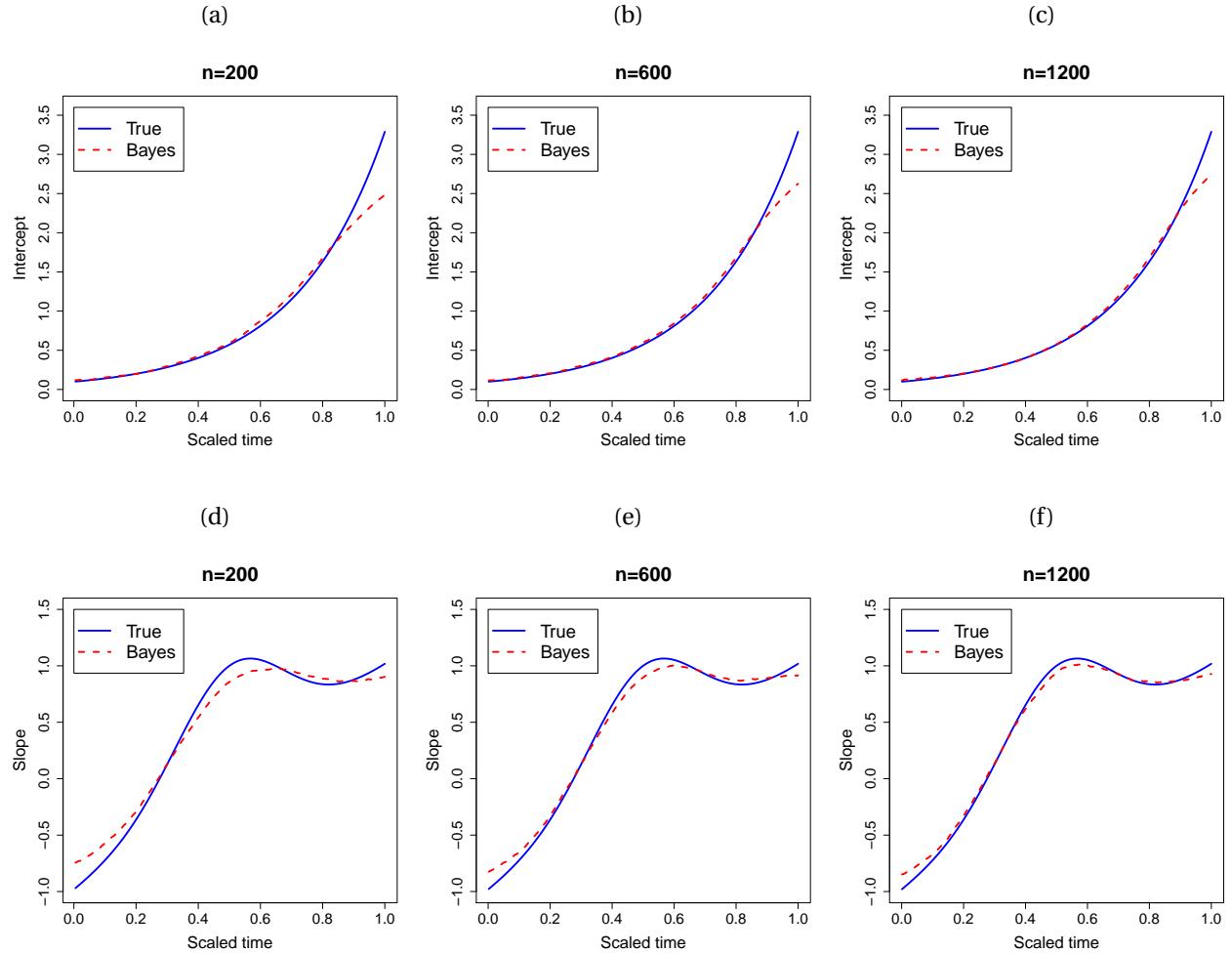


Figure 6: Median of $\hat{\beta}_1(\tau)$ and $\hat{\beta}_2(\tau)$ based on 1000 replications under DGP3 in Case 2 with sample size $n = 200, 600$ and 1200 (Graphs of $\hat{\beta}_1(\tau)$ are displayed in the first row and graphs of $\hat{\beta}_2(\tau)$ are displayed in the second row).



Appendix E: Plug-in method for bandwidth selection

In this appendix, we aim to obtain a formula for optimal bandwidth by minimizing mean squared error of local constant estimator of $\beta(\tau)$.

The local constant estimator of $\beta(\tau)$ is given by

$$\hat{\beta}(\tau; h) = \left(\sum_{t=1}^n x_t x_t^\top K_h(\tau_t - \tau) \right)^{-1} \sum_{t=1}^n x_t y_t K_h(\tau_t - \tau)$$

$$\begin{aligned}
&= \left(\sum_{t=1}^n x_t x_t^\top K_h(\tau_t - \tau) \right)^{-1} \sum_{t=1}^n x_t (x_t^\top \beta(\tau_t) + u_t) K_h(\tau_t - \tau) \\
&= \left(\sum_{t=1}^n x_t x_t^\top K_h(\tau_t - \tau) \right)^{-1} \sum_{t=1}^n x_t x_t^\top \beta(\tau_t) K_h(\tau_t - \tau) + \left(\sum_{t=1}^n x_t x_t^\top K_h(\tau_t - \tau) \right)^{-1} \sum_{t=1}^n x_t u_t K_h(\tau_t - \tau).
\end{aligned}$$

Let $p_n(\tau; h) = \frac{1}{nh} \sum_{t=1}^n x_t x_t^\top K\left(\frac{\tau_t - \tau}{h}\right)$. We then have

$$\begin{aligned}
p_n(\tau; h) &= \frac{1}{nh} \sum_{t=1}^n x_t x_t^\top K\left(\frac{\tau_t - \tau}{h}\right) = (1 + o(1)) \mathbb{E}[x_1 x_1^\top] \frac{1}{nh} \sum_{t=1}^n K\left(\frac{\tau_t - \tau}{h}\right) \\
&= (1 + o(1)) \mathbb{E}[x_1 x_1^\top] \int \frac{1}{h} K\left(\frac{u - \tau}{h}\right) du = (1 + o(1)) \mathbb{E}[x_1 x_1^\top] = \Sigma_x (1 + o(1)),
\end{aligned}$$

where $\Sigma_x = \mathbb{E}[x_1 x_1^\top]$.

Existing results from [Cai \(2007\)](#) imply

$$\text{bias}(\hat{\beta}(\tau; h)) = \mathbb{E}[\hat{\beta}(\tau; h)] - \beta(\tau) = \frac{\mu_2 h^2 \beta^{(2)}(\tau)}{2} + o(h^2), \quad (32)$$

where $\mu_2 = \int w^2 K(w) dw$.

It then follows from [Robinson \(1989\)](#) that the asymptotic variance of $\hat{\beta}(\tau; h)$ is given by

$$\text{tr}(\text{var}(\hat{\beta}(\tau; h))) = \frac{1}{nh} \sigma_u^2 \cdot \int K^2(w) dw \cdot \text{tr}(\Sigma_x^{-1}) + o\left(\frac{1}{nh}\right), \quad (33)$$

where $\sigma_u^2 = \mathbb{E}[u_1^2]$.

Therefore, based on (32) and (33), we can obtain

$$\begin{aligned}
\text{MSE}(\hat{\beta}(\tau; h)) &= \frac{h^4}{4} \mu_2^2 \|\beta^{(2)}(\tau)\|^2 + \text{tr}(\text{var}(\hat{\beta}(\tau; h))) + o(h^4 + (nh)^{-1}) \\
&= \frac{h^4}{4} \mu_2^2 \|\beta^{(2)}(\tau)\|^2 + \frac{1}{nh} \sigma_u^2 \int K^2(w) dw \text{tr}(\Sigma_x^{-1}) + o(h^4 + (nh)^{-1}) \\
&= C_1 h^4 + \frac{C_2}{nh} + o(h^4 + (nh)^{-1}),
\end{aligned}$$

where $C_1 = \frac{1}{4} \mu_2^2 \|\beta^{(2)}(\tau)\|^2$ and $C_2 = \sigma_u^2 \int K^2(w) dw \text{tr}(\Sigma_x^{-1})$.

By minimizing the leading terms of $\text{MSE}(\hat{\beta}(\tau; h))$, we can obtain the optimal bandwidth as follows:

$$h_{\text{opt}} = \left(\frac{C_2}{4C_1} \right)^{\frac{1}{5}} n^{-\frac{1}{5}}. \quad (34)$$

From (34), we can find that we need to estimate $\beta^{(2)}(\tau)$ to obtain a plug-in estimate for h . In order to do so, one may need to employ a local quadratic kernel method to estimate $\beta(\tau)$. In the simulation and empirical application, in order to avoid estimating $\beta^{(2)}(\tau)$, we use a normal reference rule (NRR), which is equivalent to h_{opt} in terms of the leading order $n^{-\frac{1}{5}}$.

References

- Cai, Z. (2007), 'Trending time-varying coefficient time series models with serially correlated errors', *Journal of Econometrics* **136**, 163–188.
- Gao, J. (2007), *Nonlinear Time Series: Semi- and Non-Parametric Methods*, Chapman & Hall/CRC, London.
- Robinson, P. M. (1989), Nonparametric estimation of time-varying parameters, in P. Hackl, ed., 'Statistical Analysis and Forecasting of Economics Structural Change', Springer, Berlin, pp. 253–264.