

Supplementary Appendix

Supplement to : Zongwu Cai, Ying Fang, Ming Lin, and Jia Su, “Inferences for a Partially Varying Coefficient Model with Endogenous Regressors”.

Lemma 1. Let $(X_1, Y_1), \dots, (X_n, Y_n)$ be i.i.d random vectors, where the Y_i ’s are scalar random variables. Assume that $\sup_x \int |y|^s f(x, y) dy < \infty$ for some $s > 2$, where $f(x, y)$ denotes the joint density of (X_i, Y_i) . Let $K(x)$ be a bounded positive function with bounded support, satisfying a Lipschitz condition, and let $K_h(x) = K(x/h)/h$. Given that $n^{2\epsilon-1}h \rightarrow \infty$ for some $\epsilon < 1 - s^{-1}$, then

$$\sup_x \left| \frac{1}{n} \sum_{k=1}^n K_h(X_k - x) Y_k - E[K_h(X_k - x) Y_k] \right| = O_p \left[\frac{\log(1/h)}{nh} \right]^{1/2}.$$

Proof of Lemma 1: The proof is given by Proposition 4 in Mack and Silverman (1982). \square

Lemma 2. Suppose that Assumptions A1-A7 hold, we have

$$[\mathbf{D}_Z^T(u)\mathbf{H}(u)\mathbf{D}_Z(u)]^{-1} \mathbf{D}_Z^T(u)\mathbf{H}(u)\mathbf{D}_X(u) = \begin{pmatrix} \Pi_X(u) & 0 \\ 0 & \Pi_X(u) \end{pmatrix} + O_p(c_n)$$

holds uniformly for $u \in \Omega$, where $c_n = \left[\frac{\log(1/h)}{nh} \right]^{1/2} + h$.

Proof of Lemma 2: Note that

$$\begin{aligned} & [\mathbf{D}_Z^T(u)\mathbf{H}(u)\mathbf{D}_Z(u)]^{-1} \mathbf{D}_Z^T(u)\mathbf{H}(u)\mathbf{D}_X(u) - \begin{pmatrix} \Pi_X(u) & 0 \\ 0 & \Pi_X(u) \end{pmatrix} \\ &= [\mathbf{D}_Z^T(u)\mathbf{H}(u)\mathbf{D}_Z(u)]^{-1} \mathbf{D}_Z^T(u)\mathbf{H}(u) \left[\mathbf{D}_X(u) - \mathbf{D}_Z(u) \begin{pmatrix} \Pi_X(u) & 0 \\ 0 & \Pi_X(u) \end{pmatrix} \right] \\ &= [\mathbf{D}_Z^T(u)\mathbf{H}(u)\mathbf{D}_Z(u)]^{-1} \\ &\quad \mathbf{D}_Z^T(u)\mathbf{H}(u) \begin{pmatrix} X_1^T - Z_1^T \Pi_X(u) & \frac{u_1-u}{h} [X_1^T - Z_1^T \Pi_X(u)] \\ \vdots & \vdots \\ X_n^T - Z_n^T \Pi_X(u) & \frac{u_n-u}{h} [X_n^T - Z_n^T \Pi_X(u)] \end{pmatrix} \end{aligned} \tag{A.1}$$

By Lemma 1, we have

$$\begin{aligned}
& \frac{1}{n} \mathbf{D}_Z^T(u) \mathbf{H}(u) \mathbf{D}_Z(u) \\
&= \frac{1}{n} \left(\begin{array}{cc} \sum_{k=1}^n Z_k Z_k^T K_h(u_k - u) & \sum_{k=1}^n Z_k Z_k^T \frac{u_k - u}{h} K_h(u_k - u) \\ \sum_{k=1}^n Z_k Z_k^T \frac{u_k - u}{h} K_h(u_k - u) & \sum_{k=1}^n Z_k Z_k^T \frac{(u_k - u)^2}{h^2} K_h(u_k - u) \end{array} \right) \\
&= \left(\begin{array}{cc} E[Z_k Z_k^T K_h(u_k - u)] & E[Z_k Z_k^T \frac{u_k - u}{h} K_h(u_k - u)] \\ E[Z_k Z_k^T \frac{u_k - u}{h} K_h(u_k - u)] & E[Z_k Z_k^T \frac{(u_k - u)^2}{h^2} K_h(u_k - u)] \end{array} \right) + O_p \left[\frac{\log(1/h)}{nh} \right]^{1/2} \\
&= \left(\begin{array}{cc} E(Z_k Z_k^T | u_k = u) f(u) + O_p(h^2) & O_p(h) \\ O_p(h) & \mu_2 E(Z_k Z_k^T | u_k = u) f(u) + O_p(h^2) \end{array} \right) + O_p \left[\frac{\log(1/h)}{nh} \right]^{1/2} \\
&= \left(\begin{array}{cc} E(Z_k Z_k^T | u_k = u) f(u) & 0 \\ 0 & \mu_2 E(Z_k Z_k^T | u_k = u) f(u) \end{array} \right) + O_p(c_n),
\end{aligned} \tag{A.2}$$

where $\mu_2 = \int t^2 K(t) dt$. Similarly, we can show that

$$\begin{aligned}
& \frac{1}{n} \mathbf{D}_Z^T(u) \mathbf{H}(u) \begin{pmatrix} X_1^T - Z_1^T \Pi_X(u) & \frac{u_1 - u}{h} [X_1^T - Z_1^T \Pi_X(u)] \\ \vdots & \vdots \\ X_n^T - Z_n^T \Pi_X(u) & \frac{u_n - u}{h} [X_n^T - Z_n^T \Pi_X(u)] \end{pmatrix} \\
&= \frac{1}{n} \mathbf{D}_Z^T(u) \mathbf{H}(u) \begin{pmatrix} v_{X,1}^T & \frac{u_1 - u}{h} v_{X,1}^T \\ \vdots & \vdots \\ v_{X,n}^T & \frac{u_n - u}{h} v_{X,n}^T \end{pmatrix} = O_P(c_n).
\end{aligned} \tag{A.3}$$

Combine (A.1), (A.2) and (A.3), we obtain that

$$[\mathbf{D}_Z^T(u) \mathbf{H}(u) \mathbf{D}_Z(u)]^{-1} \mathbf{D}_Z^T(u) \mathbf{H}(u) \mathbf{D}_X(u) - \begin{pmatrix} \Pi_X(u) & 0 \\ 0 & \Pi_X(u) \end{pmatrix} = O_p(c_n)$$

uniformly for $u \in \Omega$. This completes the proof. \square

Lemma 3 Suppose that Assumptions A1-A7 hold, then

$$\frac{1}{n} \widehat{\mathbf{D}}_X(u)^T \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) = \begin{pmatrix} \Phi(u) & 0 \\ 0 & \mu_2 \Phi(u) \end{pmatrix} + O_p(c_n)$$

uniformly for $u \in \Omega$, where $\mu_2 = \int t^2 K(t) dt$. Furthermore,

$$\left[\frac{1}{n} \widehat{\mathbf{D}}_X(u)^T \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \right]^{-1} = \begin{pmatrix} \Phi^{-1}(u) & 0 \\ 0 & \mu_2^{-1} \Phi^{-1}(u) \end{pmatrix} + O_p(c_n).$$

Proof of Lemma 3: Note that

$$\begin{aligned}
& \frac{1}{n} \widehat{\mathbf{D}}_X(u)^T \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \\
&= \mathbf{D}_X^T(u) \mathbf{H}(u) \mathbf{D}_Z(u) [\mathbf{D}_Z^T(u) \mathbf{H}(u) \mathbf{D}_Z(u)]^{-1} \left[\frac{1}{n} \mathbf{D}_Z^T(u) \mathbf{H}(u) \mathbf{D}_X(u) \right] \\
&= \mathbf{D}_X^T(u) \mathbf{H}(u) \mathbf{D}_Z(u) [\mathbf{D}_Z^T(u) \mathbf{H}(u) \mathbf{D}_Z(u)]^{-1} \\
&\quad \frac{1}{n} \begin{pmatrix} \sum_{k=1}^n Z_k X_k^T K_h(u_k - u) & \sum_{k=1}^n Z_k X_k^T \frac{u_k - u}{h} K_h(u_k - u) \\ \sum_{k=1}^n Z_k X_k^T \frac{u_k - u}{h} K_h(u_k - u) & \sum_{k=1}^n Z_k X_k^T \frac{(u_k - u)^2}{h^2} K_h(u_k - u) \end{pmatrix} \\
&= \mathbf{D}_X^T(u) \mathbf{H}(u) \mathbf{D}_Z(u) [\mathbf{D}_Z^T(u) \mathbf{H}(u) \mathbf{D}_Z(u)]^{-1} \\
&\quad \left[\begin{pmatrix} E(Z_k X_k^T \mid u_k = u) f(u) & 0 \\ 0 & \mu_2 E(Z_k X_k^T \mid u_k = u) f(u) \end{pmatrix} + O_p(c_n) \right].
\end{aligned}$$

By Lemma 2, we have

$$\begin{aligned}
& \frac{1}{n} \widehat{\mathbf{D}}_X(u)^T \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \\
&= \begin{pmatrix} \Pi_X^T(u) & 0 \\ 0 & \Pi_X^T(u) \end{pmatrix} \begin{pmatrix} E(Z_k X_k^T \mid u_k = u) f(u) & 0 \\ 0 & \mu_2 E(Z_k X_k^T \mid u_k = u) f(u) \end{pmatrix} + O_p(c_n) \\
&= \begin{pmatrix} \Pi_X^T(u) & 0 \\ 0 & \Pi_X^T(u) \end{pmatrix} \begin{pmatrix} E(Z_k Z_k^T \mid u_k = u) \Pi_X(u) f(u) & 0 \\ 0 & \mu_2 E(Z_k Z_k^T \mid u_k = u) \Pi_X(u) f(u) \end{pmatrix} + O_p(c_n) \\
&= \begin{pmatrix} \Phi(u) & 0 \\ 0 & \mu_2 \Phi(u) \end{pmatrix} + O_p(c_n).
\end{aligned}$$

This completes the proof. \square

Lemma 4 Suppose that Assumptions A1-A7 hold, then we have

$$\frac{1}{n} \widehat{\mathbf{W}}^T (\mathbf{I} - \widehat{\mathbf{S}})^T (\mathbf{I} - \widehat{\mathbf{S}}) \widehat{\mathbf{W}} \xrightarrow{p} \Sigma_1 \quad \text{and} \quad \frac{1}{n} \widehat{\mathbf{W}}^T (\mathbf{I} - \widetilde{\mathbf{S}})^T (\mathbf{I} - \widetilde{\mathbf{S}}) \mathbf{W} \xrightarrow{p} \Sigma_1,$$

where $\Sigma_1 = E(\Upsilon_k \Upsilon_k^T)$ with $\Upsilon_k = \Pi_W^T(u_k) Z_k - \Psi^T(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_k$.

Proof of Lemma 4: Note that

$$\begin{aligned}
\frac{1}{n} \widehat{\mathbf{W}}^T (\mathbf{I} - \widehat{\mathbf{S}})^T (\mathbf{I} - \widehat{\mathbf{S}}) \widehat{\mathbf{W}} &= \frac{1}{n} (\widehat{\mathbf{W}} - \widehat{\mathbf{S}} \widehat{\mathbf{W}})^T (\widehat{\mathbf{W}} - \widehat{\mathbf{S}} \widehat{\mathbf{W}}), \\
\frac{1}{n} \widehat{\mathbf{W}}^T (\mathbf{I} - \widetilde{\mathbf{S}})^T (\mathbf{I} - \widetilde{\mathbf{S}}) \mathbf{W} &= \frac{1}{n} (\widehat{\mathbf{W}} - \widetilde{\mathbf{S}} \widehat{\mathbf{W}})^T (\mathbf{W} - \widetilde{\mathbf{S}} \mathbf{W}),
\end{aligned}$$

We first consider the terms $\widehat{\mathbf{S}} \widehat{\mathbf{W}}$ and $\widetilde{\mathbf{S}} \mathbf{W}$. By the similar argument in Lemma 2, we obtain

that $\widehat{X}_k = Z_k^T \Pi_X(u_k) + O_p(c_n)$ and $\widehat{W}_k = Z_k^T \Pi_W(u_k) + O_p(c_n)$. Then by Lemma 3, we have

$$\begin{aligned}\widehat{\mathbf{S}}\widehat{\mathbf{W}} &= \begin{pmatrix} (\widehat{X}_1^T \ 0) \left[\frac{1}{n} \widehat{\mathbf{D}}_X(u_1)^T \mathbf{H}(u_1) \widehat{\mathbf{D}}_X(u_1) \right]^{-1} \frac{1}{n} \widehat{\mathbf{D}}_X(u_1)^T \mathbf{H}(u_1) \widehat{\mathbf{W}} \\ \vdots \\ (\widehat{X}_n^T \ 0) \left[\frac{1}{n} \widehat{\mathbf{D}}_X(u_n)^T \mathbf{H}(u_n) \widehat{\mathbf{D}}_X(u_n) \right]^{-1} \frac{1}{n} \widehat{\mathbf{D}}_X(u_n)^T \mathbf{H}(u_n) \widehat{\mathbf{W}} \end{pmatrix} \\ &= \begin{pmatrix} Z_1^T \Pi_X(u_1) \Phi^{-1}(u_1) \Psi(u_1) \\ \vdots \\ Z_n^T \Pi_X(u_1) \Phi^{-1}(u_n) \Psi(u_n) \end{pmatrix} + O_p(c_n),\end{aligned}$$

and

$$\begin{aligned}\widetilde{\mathbf{S}}\mathbf{W} &= \begin{pmatrix} (X_1^T \ 0) \left[\frac{1}{n} \widehat{\mathbf{D}}_X(u_1)^T \mathbf{H}(u_1) \widehat{\mathbf{D}}_X(u_1) \right]^{-1} \frac{1}{n} \widehat{\mathbf{D}}_X(u_1)^T \mathbf{H}(u_1) \mathbf{W} \\ \vdots \\ (X_n^T \ 0) \left[\frac{1}{n} \widehat{\mathbf{D}}_X(u_n)^T \mathbf{H}(u_n) \widehat{\mathbf{D}}_X(u_n) \right]^{-1} \frac{1}{n} \widehat{\mathbf{D}}_X(u_n)^T \mathbf{H}(u_n) \mathbf{W} \end{pmatrix} \\ &= \begin{pmatrix} [Z_1^T \Pi_X(u_1) + v_{X,1}^T] \Phi^{-1}(u_1) \Psi(u_1) \\ \vdots \\ [Z_n^T \Pi_X(u_1) + v_{X,n}^T] \Phi^{-1}(u_n) \Psi(u_n) \end{pmatrix} + O_p(c_n).\end{aligned}$$

Hence,

$$\begin{aligned}&\frac{1}{n} \widehat{\mathbf{W}}^T (\mathbf{I} - \widehat{\mathbf{S}})^T (\mathbf{I} - \widehat{\mathbf{S}}) \widehat{\mathbf{W}} \\ &= \frac{1}{n} (\widehat{\mathbf{W}} - \widehat{\mathbf{S}}\widehat{\mathbf{W}})^T (\widehat{\mathbf{W}} - \widehat{\mathbf{S}}\widehat{\mathbf{W}}) \\ &= \frac{1}{n} \sum_{k=1}^n [\Pi_W^T(u_k) Z_k - \Psi^T(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_k] \\ &\quad \times [Z_k^T \Pi_W(u_k) - Z_k^T \Pi_X(u_k) \Phi^{-1}(u_k) \Psi(u_k)] + O_p(c_n) \\ &\xrightarrow{p} \Sigma_1.\end{aligned}$$

Since $E(Z_k v_k^T \mid u_k) = 0$, where $v_k = (v_{X,k}^T, v_{W,k}^T)^T$, we also have

$$\begin{aligned}&\frac{1}{n} \widehat{\mathbf{W}}^T (\mathbf{I} - \widehat{\mathbf{S}})^T (\mathbf{I} - \widetilde{\mathbf{S}}) \mathbf{W} \\ &= \frac{1}{n} (\widehat{\mathbf{W}} - \widehat{\mathbf{S}}\widehat{\mathbf{W}})^T (\mathbf{W} - \widetilde{\mathbf{S}}\mathbf{W}) \\ &= \frac{1}{n} \sum_{k=1}^n [\Pi_W^T(u_k) Z_k - \Psi^T(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_k] \\ &\quad \times [Z_k^T \Pi_W(u_k) + v_{W,k}^T - Z_k^T \Pi_X(u_k) \Phi^{-1}(u_k) \Psi(u_k) - v_{X,k}^T \Phi^{-1}(u_k) \Psi(u_k)] + O_p(c_n) \\ &= \frac{1}{n} \sum_{k=1}^n [\Pi_W^T(u_k) Z_k - \Psi^T(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_k] \\ &\quad \times [Z_k^T \Pi_W(u_k) - Z_k^T \Pi_X(u_k) \Phi^{-1}(u_k) \Psi(u_k)] + O_p(c_n) \\ &\xrightarrow{p} \Sigma_1.\end{aligned}$$

We complete the proof. \square

Theorem 1. Suppose that Assumptions A1-A7 hold. When $nh^4 \rightarrow 0$, we have

$$\sqrt{n} (\tilde{\beta} - \beta) \xrightarrow{d} N(0, \Sigma_1^{-1} \Sigma_1^* \Sigma_1^{-1}),$$

where $\Sigma_1 = E(\Upsilon_k \Upsilon_k^T)$ and $\Sigma_1^* = E(\Upsilon_k \varepsilon_k^2 \Upsilon_k^T)$ with

$$\Upsilon_k = \Pi_W^T(u_k) Z_k - \Psi^T(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_k.$$

Furthermore, if ε_k is conditionally homoskedastic; that is, $E(\varepsilon_k^2 | u_k, Z_k) = \sigma_\varepsilon^2$, then, we have

$$\sqrt{n} (\tilde{\beta} - \beta) \xrightarrow{d} N(0, \sigma_\varepsilon^2 \Sigma_1^{-1})$$

since $\Sigma_1^* = \Sigma_1 \sigma_\varepsilon^2$.

Proof of Theorem 1: The estimator $\tilde{\beta}$ is given by

$$\begin{aligned} \tilde{\beta} &= \{\widehat{\mathbf{W}}^T(\mathbf{I} - \widehat{\mathbf{S}})^T(\mathbf{I} - \widetilde{\mathbf{S}})\mathbf{W}\}^{-1} \widehat{\mathbf{W}}^T(\mathbf{I} - \widehat{\mathbf{S}})^T(\mathbf{I} - \widetilde{\mathbf{S}})\mathbf{Y} \\ &= \{\widehat{\mathbf{W}}^T(\mathbf{I} - \widehat{\mathbf{S}})^T(\mathbf{I} - \widetilde{\mathbf{S}})\mathbf{W}\}^{-1} \widehat{\mathbf{W}}^T(\mathbf{I} - \widehat{\mathbf{S}})^T(\mathbf{I} - \widetilde{\mathbf{S}})(\mathbf{W}\beta + \mathbf{M} + \boldsymbol{\varepsilon}), \end{aligned}$$

then

$$\begin{aligned} \sqrt{n} (\tilde{\beta} - \beta) &= \left\{ \frac{1}{n} \widehat{\mathbf{W}}^T(\mathbf{I} - \widehat{\mathbf{S}})^T(\mathbf{I} - \widetilde{\mathbf{S}})\mathbf{W} \right\}^{-1} \\ &\quad \left[\frac{1}{\sqrt{n}} \widehat{\mathbf{W}}^T(\mathbf{I} - \widehat{\mathbf{S}})^T(\mathbf{I} - \widetilde{\mathbf{S}})\mathbf{M} + \frac{1}{\sqrt{n}} \widehat{\mathbf{W}}^T(\mathbf{I} - \widehat{\mathbf{S}})^T(\mathbf{I} - \widetilde{\mathbf{S}})\boldsymbol{\varepsilon} \right]. \end{aligned} \tag{A.4}$$

The first term converges to Σ_1^{-1} by Lemma 4, we now consider the second term $\frac{1}{\sqrt{n}} \widehat{\mathbf{W}}^T(\mathbf{I} - \widehat{\mathbf{S}})^T(\mathbf{I} - \widetilde{\mathbf{S}})\mathbf{M}$ and the last term $\frac{1}{\sqrt{n}} \widehat{\mathbf{W}}^T(\mathbf{I} - \widehat{\mathbf{S}})^T(\mathbf{I} - \widetilde{\mathbf{S}})\boldsymbol{\varepsilon}$.

First, note that

$$\frac{1}{\sqrt{n}} \widehat{\mathbf{W}}^T(\mathbf{I} - \widehat{\mathbf{S}})^T(\mathbf{I} - \widetilde{\mathbf{S}})\mathbf{M} = \frac{1}{\sqrt{n}} (\widehat{\mathbf{W}} - \widehat{\mathbf{S}}\widehat{\mathbf{W}})^T(\mathbf{M} - \widetilde{\mathbf{S}}\mathbf{M}).$$

Define $r(u_k) = \bar{A}(u_k) - A(u_k)$, where $\bar{A}(u_k)$ is the vector consisting of the first p rows of $[\widehat{\mathbf{D}}_X(u_k)^T \mathbf{H}(u_k) \widehat{\mathbf{D}}_X(u_k)]^{-1} \widehat{\mathbf{D}}_X(u_k)^T \mathbf{H}(u_k) \mathbf{M}$. Then

$$\begin{aligned} \widetilde{\mathbf{S}}\mathbf{M} &= \begin{pmatrix} (X_1^T \mathbf{0}_{1 \times p}) [\widehat{\mathbf{D}}_X(u_1)^T \mathbf{H}(u_1) \widehat{\mathbf{D}}_X(u_1)]^{-1} \widehat{\mathbf{D}}_X(u_1)^T \mathbf{H}(u_1) \mathbf{M} \\ \vdots \\ (X_n^T \mathbf{0}_{1 \times p}) [\widehat{\mathbf{D}}_X(u_n)^T \mathbf{H}(u_n) \widehat{\mathbf{D}}_X(u_n)]^{-1} \widehat{\mathbf{D}}_X(u_n)^T \mathbf{H}(u_n) \mathbf{M} \end{pmatrix} \\ &= \begin{pmatrix} X_1^T [A(u_1) + r(u_1)] \\ \vdots \\ X_n^T [A(u_n) + r(u_n)] \end{pmatrix}. \end{aligned}$$

By the similar argument in Lemma 4, it can be shown that $r(u_k) = O_p(c_n)$ uniformly for $u_k \in \Omega$. Note that $\sqrt{n}O_p(c_n^2) = o_p(1)$, we have

$$\begin{aligned}
& \frac{1}{\sqrt{n}} \widehat{\mathbf{W}}^T (\mathbf{I} - \widehat{\mathbf{S}})^T (\mathbf{I} - \widetilde{\mathbf{S}}) \mathbf{M} \\
= & \frac{1}{\sqrt{n}} (\widehat{\mathbf{W}} - \widehat{\mathbf{S}} \widehat{\mathbf{W}})^T (\mathbf{M} - \widetilde{\mathbf{S}} \mathbf{M}) \\
= & \frac{1}{\sqrt{n}} \sum_{k=1}^n [\Pi_W^T(u_k) Z_k - \Psi(u_k)^T \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_k + O_p(c_n)] \{X_k^T A(u_k) - X_k^T [A(u_k) + r(u_k)]\} \\
= & \frac{1}{\sqrt{n}} \sum_{k=1}^n [\Pi_W^T(u_k) Z_k - \Psi(u_k)^T \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_k + O_p(c_n)] X_k^T r(u_k) \\
= & \frac{1}{\sqrt{n}} \sum_{k=1}^n [\Pi_W^T(u_k) Z_k - \Psi(u_k)^T \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_k] X_k^T r(u_k) + o_p(1) \\
= & \frac{1}{\sqrt{n}} \sum_{k=1}^n L_k(u_k) r(u_k) + o_p(1),
\end{aligned}$$

where

$$L_k(u_k) = [\Pi_W^T(u_k) Z_k - \Psi(u_k)^T \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_k] X_k^T.$$

Let $r_{-k}(u_k)$ be the leave-one-out version of $r(u_k)$ without using the k -th observation, that is, replace all $\mathbf{H}(u_k)$'s in $r(u_k)$ by $\mathbf{H}_{-k}(u_k) = \text{diag}(K_h(u_1 - u_k), \dots, K_h(u_{k-1} - u_k), 0, K_h(u_{k+1} - u_k), \dots, K_h(u_n - u_k))$. It can be shown that $r(u_k) - r_{-k}(u_k) = O_p(\frac{1}{nh})$. Since $nh^2 \rightarrow \infty$, we have

$$\frac{1}{\sqrt{n}} \widehat{\mathbf{W}}^T (\mathbf{I} - \widehat{\mathbf{S}})^T (\mathbf{I} - \widetilde{\mathbf{S}}) \mathbf{M} = \frac{1}{\sqrt{n}} \sum_{k=1}^n L_k(u_k) r_{-k}(u_k) + o_p(1). \quad (\text{A.5})$$

Conditional on u_k , $L_k(u_k)$ and $r_{-k}(u_k)$ are independent, and $E[L_k(u_k) | u_k] = 0$, so

$$E \left[\frac{1}{\sqrt{n}} \sum_{k=1}^n L_k(u_k) r_{-k}(u_k) \right] = \frac{1}{\sqrt{n}} \sum_{k=1}^n E \{E[L_k(u_k) | u_k] E[r_{-k}(u_k) | u_k]\} = 0. \quad (\text{A.6})$$

and

$$\begin{aligned}
& \text{Var} \left[\frac{1}{\sqrt{n}} \sum_{k=1}^n L_k(u_k) r_{-k}(u_k) \right] \\
= & \frac{1}{n} \sum_{k=1}^n \text{Var}[L_k(u_k) r_{-k}(u_k)] + \frac{1}{n} \sum_{j \neq k} \text{Cov}[L_j(u_j) r_{-j}(u_j), L_k(u_k) r_{-k}(u_k)] \\
= & \frac{1}{n} \sum_{j \neq k} E[L_j(u_j) r_{-j}(u_j) L_k(u_k) r_{-k}(u_k)] + o(1).
\end{aligned}$$

We further let $r_{-jk}(u_k)$ be the jackknife estimator of $r(u_k)$ without using the j, k -th observations, then for $j \neq k$,

$$\begin{aligned}
& E [L_j(u_j)r_{-j}(u_j)L_k(u_k)r_{-k}(u_k) | u_j, u_k] \\
= & E [L_j(u_j)r_{-jk}(u_j)L_k(u_k)r_{-jk}(u_k) | u_j, u_k] \\
& + E \{L_j(u_j)[r_{-j}(u_j) - r_{-jk}(u_j)]L_k(u_k)r_{-jk}(u_k) | u_j, u_k\} \\
& + E \{L_j(u_j)r_{-jk}(u_j)L_k(u_k)[r_{-k}(u_k) - r_{-jk}(u_k)] | u_j, u_k\} \\
& + E \{L_j(u_j)[r_{-j}(u_j) - r_{-jk}(u_j)]L_k(u_k)[r_{-k}(u_k) - r_{-jk}(u_k)] | u_j, u_k\} \\
= & 0 + 0 + 0 + O\left(\frac{1}{n^2 h^2}\right).
\end{aligned}$$

Therefore, we have

$$\text{Var} \left[\frac{1}{\sqrt{n}} \sum_{k=1}^n L_k(u_k)r_{-k}(u_k) \right] = \frac{1}{n} \cdot n^2 \cdot O\left(\frac{1}{n^2 h^2}\right) + o(1) \rightarrow 0. \quad (\text{A.7})$$

From (A.5), (A.6) and (A.7), we can obtain that

$$\frac{1}{\sqrt{n}} \widehat{\mathbf{W}}^T (\mathbf{I} - \widehat{\mathbf{S}})^T (\mathbf{I} - \widetilde{\mathbf{S}}) \mathbf{M} = o_p(1). \quad (\text{A.8})$$

Next, note that

$$\begin{aligned}
\widetilde{\mathbf{S}}\boldsymbol{\varepsilon} &= \begin{pmatrix} (X_1^T \ 0) [\widehat{\mathbf{D}}_X(u_1)^T \mathbf{H}(u_1) \widehat{\mathbf{D}}_X(u_1)]^{-1} \widehat{\mathbf{D}}_X(u_1)^T \mathbf{H}(u_1) \boldsymbol{\varepsilon} \\ \vdots \\ (X_n^T \ 0) [\widehat{\mathbf{D}}_X(u_n)^T \mathbf{H}(u_n) \widehat{\mathbf{D}}_X(u_n)]^{-1} \widehat{\mathbf{D}}_X(u_n)^T \mathbf{H}(u_n) \boldsymbol{\varepsilon} \end{pmatrix} \\
&= \begin{pmatrix} X_1^T O_p(c_n) \\ \vdots \\ X_n^T O_p(c_n) \end{pmatrix},
\end{aligned}$$

we have

$$\begin{aligned}
& \frac{1}{\sqrt{n}} \widehat{\mathbf{W}}^T (\mathbf{I} - \widehat{\mathbf{S}})^T (\mathbf{I} - \widetilde{\mathbf{S}}) \boldsymbol{\varepsilon} \\
= & \frac{1}{\sqrt{n}} \sum_{k=1}^n [\Pi_W^T(u_k) Z_k - \boldsymbol{\Psi}(u_k)^T \boldsymbol{\Phi}^{-1}(u_k) \Pi_X^T(u_k) Z_k + O_p(c_n)] [\varepsilon_k - X_k^T O_p(c_n)] \\
= & \frac{1}{\sqrt{n}} \sum_{k=1}^n [\Pi_W^T(u_k) Z_k - \boldsymbol{\Psi}^T(u_k) \boldsymbol{\Phi}^{-1}(u_k) \Pi_X^T(u_k) Z_k] \varepsilon_k + \frac{1}{\sqrt{n}} \sum_{k=1}^n \varepsilon_k O_p(c_n) \\
& - \frac{1}{\sqrt{n}} \sum_{k=1}^n [\Pi_W^T(u_k) Z_k - \boldsymbol{\Psi}^T(u_k) \boldsymbol{\Phi}^{-1}(u_k) \Pi_X^T(u_k) Z_k] X_k O_p(c_n) + o_p(1).
\end{aligned}$$

Since $E(\varepsilon_k \mid u_k) = 0$ and $E \left\{ [\Pi_W^T(u_k)Z_k - \Psi^T(u_k)\Phi^{-1}(u_k)\Pi_X^T(u_k)Z_k] X_k \mid u_k \right\} = 0$, by the similar argument of obtaining (A.8), we have

$$\begin{aligned} & \frac{1}{\sqrt{n}} \widehat{\mathbf{W}}^T (\mathbf{I} - \widehat{\mathbf{S}})^T (\mathbf{I} - \widetilde{\mathbf{S}}) \varepsilon \\ &= \frac{1}{\sqrt{n}} \sum_{k=1}^n [\Pi_W^T(u_k)Z_k - \Psi^T(u_k)\Phi^{-1}(u_k)\Pi_X^T(u_k)Z_k] \varepsilon_k + o_p(1) + o_p(1) + o_p(1) \\ &\xrightarrow{d} N(0, \Sigma_1^*) . \end{aligned} \quad (\text{A.9})$$

By Lemma 4 and (A.4), (A.8), (A.9), we have

$$\sqrt{n} (\widetilde{\beta} - \beta) \xrightarrow{d} N(0, \Sigma_1^{-1} \Sigma_1^* \Sigma_1^{-1}) .$$

This completes the proof. \square

Theorem 2. Suppose that $\bar{\beta}$ is a \sqrt{n} -consistent estimate of β and under Assumptions A1-A7, we have

$$\sqrt{nh} \left[\widetilde{A}(u) - A(u) - \frac{1}{2} h^2 \mu_2 A''(u) (1 + o_p(1)) \right] \xrightarrow{d} N(0, \Sigma_2) ,$$

where $\mu_2 = \int t^2 K(t) dt$ and $\Sigma_2 = \nu_0 \Phi^{-1}(u) \Lambda(u) \Phi^{-1}(u)$ with $\nu_0 = \int K^2(t) dt$. If ε_k is conditionally homoskedastic, then, the asymptotic variance reduces to $\Sigma_2 = \nu_0 \sigma_\varepsilon^2 \Phi^{-1}(u)$ since $\Lambda(u) = \sigma_\varepsilon^2 \Phi(u)$. Furthermore, if $nh^5 \rightarrow 0$, we have

$$\sqrt{nh} [\widetilde{A}(u) - A(u)] \xrightarrow{d} N(0, \Sigma_2) . \quad (\text{A.10})$$

Proof of Theorem 2: Note that $\widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) = \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \mathbf{D}_X(u)$, we have

$$\begin{aligned} & \sqrt{nh} \left[\begin{pmatrix} \widetilde{A}(u) \\ h\widetilde{A}'(u) \end{pmatrix} - \begin{pmatrix} A(u) \\ hA'(u) \end{pmatrix} \right] \\ &= \sqrt{nh} \left\{ [\widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u)]^{-1} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) (\mathbf{Y} - \mathbf{W}\bar{\beta}) - \begin{pmatrix} A(u) \\ hA'(u) \end{pmatrix} \right\} \\ &= \sqrt{nh} \left[\widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \right]^{-1} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \left[\mathbf{W}(\beta - \bar{\beta}) + \mathbf{M} + \boldsymbol{\varepsilon} - \widehat{\mathbf{D}}_X(u) \begin{pmatrix} A(u) \\ hA'(u) \end{pmatrix} \right] \\ &= \sqrt{nh} \left[\widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \right]^{-1} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \left[\mathbf{W}(\beta - \bar{\beta}) + \mathbf{M} + \boldsymbol{\varepsilon} - \mathbf{D}_X(u) \begin{pmatrix} A(u) \\ hA'(u) \end{pmatrix} \right] \\ &= \left[\frac{1}{n} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \right]^{-1} \\ & \quad \sqrt{\frac{h}{n}} \left\{ \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \mathbf{W}(\beta - \bar{\beta}) + \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \left[\mathbf{M} - \mathbf{D}_X(u) \begin{pmatrix} A(u) \\ hA'(u) \end{pmatrix} \right] + \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \boldsymbol{\varepsilon} \right\} \\ &\triangleq \left[\frac{1}{n} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \right]^{-1} (J_1 + J_2 + J_3). \end{aligned}$$

We first consider the term J_1 . By Lemma 3 and the square root consistency of $\bar{\beta}$, we have

$$\begin{aligned} J_1 &= \sqrt{\frac{h}{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \mathbf{W}(\beta - \bar{\beta}) \\ &= \sqrt{h} \left[\frac{1}{n} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \mathbf{W} \right] [\sqrt{n}(\beta - \bar{\beta})] = o_p(1). \end{aligned} \quad (\text{A.11})$$

We now move to the term J_2 . By Lemma 2,

$$\begin{aligned} J_2 &= \sqrt{\frac{h}{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \left[\begin{pmatrix} X_1^T A(u_1) \\ \vdots \\ X_n^T A(u_n) \end{pmatrix} - \mathbf{D}_X(u) \begin{pmatrix} A(u) \\ hA'(u) \end{pmatrix} \right] \\ &= \sqrt{\frac{h}{n}} \left[\begin{pmatrix} \Pi_X^T(u) & 0 \\ 0 & \Pi_X^T(u) \end{pmatrix} + o_p(1) \right] \mathbf{D}_Z^T(u) \mathbf{H}(u) \\ &\quad \left[\begin{pmatrix} X_1^T A(u_1) \\ \vdots \\ X_n^T A(u_n) \end{pmatrix} - \begin{pmatrix} X_1^T A(u) + X_1^T A'(u)(u_1 - u) \\ \vdots \\ X_1^T A(u) + X_1^T A'(u)(u_n - u) \end{pmatrix} \right] \\ &= \sqrt{nh} \left[\begin{pmatrix} \Pi_X^T(u) & 0 \\ 0 & \Pi_X^T(u) \end{pmatrix} + o_p(1) \right] \frac{1}{n} \mathbf{D}_Z^T(u) \mathbf{H}(u) \left[\begin{pmatrix} X_1^T A''(\xi_1)(u_1 - u)^2/2 \\ \vdots \\ X_1^T A''(\xi_n)(u_n - u)^2/2 \end{pmatrix} \right] \\ &= \sqrt{nh} \begin{pmatrix} \frac{1}{2}h^2 \Phi(u) [\mu_2 A''(u) + o_p(1)] \\ o_p(h^2) \end{pmatrix}. \end{aligned} \quad (\text{A.12})$$

Here ξ_k is a point between u_k and u .

Finally, we consider the term J_3 . Note that

$$\begin{aligned} J_3 &= \sqrt{\frac{h}{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \boldsymbol{\varepsilon} \\ &= \left[\begin{pmatrix} \Pi_X^T(u) & 0 \\ 0 & \Pi_X^T(u) \end{pmatrix} + o_p(1) \right] \sqrt{\frac{h}{n}} \mathbf{D}_Z^T(u) \mathbf{H}(u) \boldsymbol{\varepsilon} \\ &= \begin{pmatrix} \Pi_X^T(u) & 0 \\ 0 & \Pi_X^T(u) \end{pmatrix} \left(\frac{1}{\sqrt{n}} \sum_{k=1}^n \sqrt{h} Z_k \varepsilon_k K_h(u_k - u) \right) + o_p(1) \end{aligned}$$

Here, we only focus on the variance of $\widehat{A}(u)$, then we have

$$\begin{aligned} \text{Var} \left[\sqrt{h} Z_k \varepsilon_k K_h(u_k - u) \right] &= E [h Z_k \varepsilon_k^2 Z_k^T K_h^2(u_k - u)] \\ &= \int K^2(t) dt E [Z_k \varepsilon_k^2 Z_k^T \mid u_k = u] f(u) + o_p(1) \\ &= \nu_0 E [Z_k \varepsilon_k^2 Z_k^T \mid u_k = u] f(u) + o_p(1). \end{aligned}$$

Hence, we have

$$J_3 \xrightarrow{d} N(0, \nu_0 \boldsymbol{\Lambda}(u)). \quad (\text{A.13})$$

By Lemma 3, we have

$$\left[\frac{1}{n} \widehat{\mathbf{D}}_X(u)^T \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \right]^{-1} \xrightarrow{p} \begin{pmatrix} \Phi^{-1}(u) & 0 \\ 0 & \mu_2^{-1} \Phi^{-1}(u) \end{pmatrix},$$

together with (A.11), (A.12) and (A.13), we have

$$\sqrt{nh} \left[\widetilde{A}(u) - A(u) - \frac{1}{2} h^2 \mu_2 A''(u) (1 + o_p(1)) \right] \xrightarrow{d} N(0, \nu_0 \Phi^{-1}(u) \Lambda(u) \Phi^{-1}(u)).$$

This completes the proof. \square

Lemma 5 Suppose Assumptions A1-A7 hold, then under the null hypothesis,

$$\frac{1}{n} \text{RSS}_1 \xrightarrow{p} E \left\{ [\varepsilon_k + v_{X,k}^T A_0(u_k; \theta)]^2 f(u_k) \right\}.$$

Proof of Lemma 5: We have

$$\begin{aligned} \frac{1}{n} \text{RSS}_1 &= \frac{1}{n} \sum_{k=1}^n \frac{1}{n} \left[\mathbf{Y} - \widehat{\mathbf{D}}_X(u_k) \begin{pmatrix} \widetilde{A}(u_k) \\ h\widetilde{A}'(u_k) \end{pmatrix} - \mathbf{W}\bar{\beta} \right]^T \\ &\quad \mathbf{H}(u_k) \left[\mathbf{Y} - \widehat{\mathbf{D}}_X(u_k) \begin{pmatrix} \widetilde{A}(u_k) \\ h\widetilde{A}'(u_k) \end{pmatrix} - \mathbf{W}\bar{\beta} \right] \\ &= \frac{1}{n} \sum_{k=1}^n \frac{1}{n} \left[\mathbf{Y} - \widehat{\mathbf{D}}_X(u_k) \begin{pmatrix} A_0(u_k; \theta) \\ hA'_0(u_k; \theta) \end{pmatrix} - \mathbf{W}\beta \right]^T \\ &\quad \mathbf{H}(u_k) \left[\mathbf{Y} - \widehat{\mathbf{D}}_X(u_k) \begin{pmatrix} A_0(u_k; \theta) \\ hA'_0(u_k; \theta) \end{pmatrix} - \mathbf{W}\beta \right] + o_p(1). \end{aligned}$$

According to Lemma 2, we have

$$\begin{aligned} \frac{1}{n} \text{RSS}_1 &= \frac{1}{n} \sum_{k=1}^n \frac{1}{n} \left[\mathbf{Y} - \mathbf{D}_Z(u_k) \begin{pmatrix} \Pi(u_k) A_0(u_k; \theta) \\ \Pi(u_k) hA'_0(u_k; \theta) \end{pmatrix} - \mathbf{W}\beta \right]^T \\ &\quad \mathbf{H}(u_k) \left[\mathbf{Y} - \mathbf{D}_Z(u_k) \begin{pmatrix} \Pi(u_k) A_0(u_k; \theta) \\ \Pi(u_k) hA'_0(u_k; \theta) \end{pmatrix} - \mathbf{W}\beta \right] + o_p(1) \\ &= \frac{1}{n} \sum_{k=1}^n \frac{1}{n} \begin{pmatrix} \varepsilon_1 + v_{X,1}^T A_0(u_1, \theta) \\ \vdots \\ \varepsilon_n + v_{X,n}^T A_0(u_n, \theta) \end{pmatrix}^T \mathbf{H}(u_k) \begin{pmatrix} \varepsilon_1 + v_{X,1}^T A_0(u_1, \theta) \\ \vdots \\ \varepsilon_n + v_{X,n}^T A_0(u_n, \theta) \end{pmatrix} + o_p(1) \\ &= \frac{1}{n} \sum_{k=1}^n E \left\{ [\varepsilon_k + v_{X,k}^T A_0(u_k; \theta)]^2 \mid u_k \right\} f(u_k) + o_p(1) \\ &= E \left\{ [\varepsilon_k + v_{X,k}^T A_0(u_k; \theta)]^2 f(u_k) \right\} + o_p(1). \end{aligned}$$

The proof is complete. \square

Lemma 6 Let $\bar{\theta}$ and $\bar{\beta}$ be \sqrt{n} -consistent estimates of θ and β under the null hypothesis, respectively. Define

$$S_0(u) = \left[\mathbf{Y} - \widehat{\mathbf{D}}_X(u) \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} - \mathbf{W}\bar{\beta} \right]^T \mathbf{H}(u) \left[\mathbf{Y} - \widehat{\mathbf{D}}_X(u) \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} - \mathbf{W}\bar{\beta} \right],$$

$$S_1(u) = \left[\mathbf{Y} - \widehat{\mathbf{D}}_X(u) \begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \mathbf{W}\bar{\beta} \right]^T \mathbf{H}(u) \left[\mathbf{Y} - \widehat{\mathbf{D}}_X(u) \begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \mathbf{W}\bar{\beta} \right],$$

where

$$\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} = \left[\widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \right]^{-1} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) (\mathbf{Y} - \mathbf{W}\bar{\beta}).$$

Suppose that Assumptions A1-A7 hold and $nh^4 \rightarrow 0$, then under the null hypothesis,

$$S_0(u) - S_1(u) = P(u) + T_1(u) + T_2(u) + o_p(1/\sqrt{h}),$$

where

$$P(u) = \frac{1}{n} \boldsymbol{\varepsilon}^T \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \left[\frac{1}{n} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \right]^{-1} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \boldsymbol{\varepsilon},$$

$$T_1(u) = \frac{1}{n} (\beta - \bar{\beta})^T \mathbf{W}^T \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \left[\frac{1}{n} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \right]^{-1} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \boldsymbol{\varepsilon},$$

$$T_2(u) = \frac{1}{n} (\theta - \bar{\theta})^T \frac{\partial A_0^T(u; \theta)}{\partial \theta} \mathbf{X}^T \mathbf{H}(u) \widehat{\mathbf{D}}_X(u)$$

$$\left[\frac{1}{n} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \right]^{-1} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \boldsymbol{\varepsilon}.$$

Proof of Lemma 6: Note that we have

$$S_0(u)$$

$$= \left\{ \mathbf{Y} - \widehat{\mathbf{D}}_X(u) \begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \mathbf{W}\bar{\beta} + \widehat{\mathbf{D}}_X(u) \left[\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right] \right\}^T$$

$$\mathbf{H}(u) \left\{ \mathbf{Y} - \widehat{\mathbf{D}}_X(u) \begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \mathbf{W}\bar{\beta} + \widehat{\mathbf{D}}_X(u) \left[\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right] \right\}$$

$$= S_1(u) + 2 \left[\mathbf{Y} - \widehat{\mathbf{D}}_X(u) \begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \mathbf{W}\bar{\beta} \right]^T \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \left[\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right]$$

$$+ \left[\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right]^T \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \left[\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right].$$

Because

$$\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} = \left[\widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \widehat{\mathbf{D}}_X(u) \right]^{-1} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) (\mathbf{Y} - \mathbf{W}\bar{\beta}),$$

so the cross term

$$\begin{aligned}
& \left[\mathbf{Y} - \widehat{\mathbf{D}}_X(u) \begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \mathbf{W}\bar{\beta} \right]^T \mathbf{H}(u)\widehat{\mathbf{D}}_X(u) \left[\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right] \\
&= (\mathbf{Y} - \mathbf{W}\bar{\beta})^T \left\{ \mathbf{I}_n - \mathbf{H}(u)\widehat{\mathbf{D}}_X(u) \left[\widehat{\mathbf{D}}_X^T(u)\mathbf{H}(u)\widehat{\mathbf{D}}_X(u) \right]^{-1} \widehat{\mathbf{D}}_X^T(u) \right\} \\
&\quad \mathbf{H}(u)\widehat{\mathbf{D}}_X(u) \left[\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right] \\
&= (\mathbf{Y} - \mathbf{W}\bar{\beta})^T \left[\mathbf{H}(u)\widehat{\mathbf{D}}_X(u) - \mathbf{H}(u)\widehat{\mathbf{D}}_X(u) \right] \left[\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right] \\
&= 0.
\end{aligned}$$

Furthermore, because $\widehat{\mathbf{D}}_X^T(u)\mathbf{H}(u)\widehat{\mathbf{D}}_X(u) = \widehat{\mathbf{D}}_X^T(u)\mathbf{H}(u)\mathbf{D}_X(u)$, we have

$$\begin{aligned}
& S_0(u) - S_1(u) \\
&= \left[\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right]^T \widehat{\mathbf{D}}_X^T(u)\mathbf{H}(u)\widehat{\mathbf{D}}_X(u) \left[\begin{pmatrix} \tilde{A}(u) \\ h\tilde{A}'(u) \end{pmatrix} - \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right] \\
&= \left[\mathbf{Y} - \mathbf{W}\bar{\beta} - \widehat{\mathbf{D}}_X(u) \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right]^T \mathbf{H}(u)\widehat{\mathbf{D}}_X(u) \\
&\quad \left[\widehat{\mathbf{D}}_X^T(u)\mathbf{H}(u)\widehat{\mathbf{D}}_X(u) \right]^{-1} \widehat{\mathbf{D}}_X^T(u)\mathbf{H}(u) \left[\mathbf{Y} - \mathbf{W}\bar{\beta} - \widehat{\mathbf{D}}_X(u) \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right] \\
&= \frac{1}{\sqrt{n}} \left[\mathbf{Y} - \mathbf{W}\bar{\beta} - \mathbf{D}_X(u) \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right]^T \mathbf{H}(u)\widehat{\mathbf{D}}_X(u) \\
&\quad \left[\frac{1}{n} \widehat{\mathbf{D}}_X^T(u)\mathbf{H}(u)\widehat{\mathbf{D}}_X(u) \right]^{-1} \frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u)\mathbf{H}(u) \left[\mathbf{Y} - \mathbf{W}\bar{\beta} - \mathbf{D}_X(u) \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right].
\end{aligned}$$

Note that $\bar{\beta}$ and $\bar{\theta}$ are \sqrt{n} -consistent estimates. When $nh^4 \rightarrow 0$, we have

$$\begin{aligned}
& \frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \left[\mathbf{Y} - \mathbf{W}\bar{\beta} - \mathbf{D}_X(u) \begin{pmatrix} A_0(u; \bar{\theta}) \\ hA'_0(u; \bar{\theta}) \end{pmatrix} \right] \\
= & \frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \left[\mathbf{Y} - \mathbf{W}\beta - \begin{pmatrix} X_1^T A_0(u_1; \theta) \\ \vdots \\ X_n^T A_0(u_n; \theta) \end{pmatrix} \right] + \frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \mathbf{W}(\beta - \bar{\beta}) \\
& + \frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \begin{pmatrix} X_1^T [A_0(u_1; \theta) - A_0(u; \bar{\theta}) - A'_0(u; \bar{\theta})(u_1 - u)] \\ \vdots \\ X_n^T [A_0(u_n; \theta) - A_0(u; \bar{\theta}) - A'_0(u; \bar{\theta})(u_n - u)] \end{pmatrix} \\
= & \frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \boldsymbol{\varepsilon} + \frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \mathbf{W}(\beta - \bar{\beta}) \\
& + \frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \mathbf{X} \frac{\partial A_0(u; \theta)}{\partial \theta} (\theta - \bar{\theta}) + O_p \left[\sqrt{n} \left(\frac{1}{n} + \frac{1}{\sqrt{n}} h + h^2 \right) \right] \\
= & \frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \boldsymbol{\varepsilon} + \frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \mathbf{W}(\beta - \bar{\beta}) \\
& + \frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \mathbf{X} \frac{\partial A_0(u; \theta)}{\partial \theta} (\theta - \bar{\theta}) + o_p(1).
\end{aligned}$$

Since $\frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \mathbf{W}(\beta - \bar{\beta}) = O_p(1)$ and $\frac{1}{\sqrt{n}} \widehat{\mathbf{D}}_X^T(u) \mathbf{H}(u) \mathbf{X} \frac{\partial A_0(u; \theta)}{\partial \theta} (\theta - \bar{\theta}) = O_p(1)$, then

$$\begin{aligned}
S_0(u) - S_1(u) &= P(u) + T_1(u) + T_2(u) + O_p(1) \\
&= P(u) + T_1(u) + T_2(u) + o_p(1/\sqrt{h}).
\end{aligned}$$

This completes the proof.

Theorem 3. Suppose that Assumptions A1-A7 hold, when $nh^4 \rightarrow 0$, under the null hypothesis, we have

$$\sigma_n^{-1} \{ \lambda_n - \mu_n \} \xrightarrow{d} N(0, 1),$$

where

$$\begin{aligned}
\mu_n &= \frac{(\nu_0 + \mu_2^{-1} \nu_2) E \left\{ \text{Trace} \{ \Phi^{-1}(u_k) \Lambda(u_k) \} \right\}}{h E \left\{ [\varepsilon_k + v_{X,k}^T A_0(u_k; \theta)]^2 f(u_k) \right\}} \\
\text{and } \sigma_n^2 &= \frac{2 \int g^2(t) dt E \left\{ \text{Trace} \{ \Phi^{-1}(u_k) \Lambda(u_k) \Phi^{-1}(u_k) \Lambda(u_k) f(u_k) \} \right\}}{h E^2 \left\{ [\varepsilon_k + v_{X,k}^T A_0(u_k; \theta)]^2 f(u_k) \right\}}
\end{aligned}$$

with $\nu_0 = \int K^2(t) dt$, $\nu_2 = \int t^2 K^2(t) dt$, $\mu_2 = \int t^2 K(t) dt$ and $g(t) = \int K(s) K(t+s) ds + \mu_2^{-1} \int s(t+s) K(s) K(t+s) ds$. Furthermore, if ε_k is conditionally homoskedastic, then,

$\Phi^{-1}(u_k)\Lambda(u_k) = \sigma_\varepsilon^2 \mathbf{I}_p$ so that

$$\mu_n = \frac{(\nu_0 + \mu_2^{-1}\nu_2)p\sigma_\varepsilon^2}{h E \left\{ [\varepsilon_k + v_{X,k}^T A_0(u_k; \theta)]^2 f(u_k) \right\}}$$

$$\text{and } \sigma_n^2 = \frac{2 \int g^2(t) dt p \sigma_\varepsilon^4 E[f(u_k)]}{h E^2 \left\{ [\varepsilon_k + v_{X,k}^T A_0(u_k; \theta)]^2 f(u_k) \right\}}.$$

Proof of Theorem 3: Note that

$$\begin{aligned} \text{RSS}_0 - \text{RSS}_1 &= \frac{1}{n} \sum_{k=1}^n [S_0(u_k) - S_1(u_k)] \\ &= \mathbf{P} + \mathbf{T}_1 + \mathbf{T}_2 + o_p(1/\sqrt{h}), \end{aligned}$$

where $\mathbf{P} = \frac{1}{n} \sum_{k=1}^n P(u_k)$, $\mathbf{T}_1 = \frac{1}{n} \sum_{k=1}^n T_1(u_k)$ and $\mathbf{T}_2 = \frac{1}{n} \sum_{k=1}^n T_2(u_k)$.

By Lemma 3, we have

$$\begin{aligned} \mathbf{T}_1 &= \frac{1}{\sqrt{h}} \sqrt{n}(\beta_0 - \bar{\beta})^T \frac{1}{n} \sum_{k=1}^n \left[\frac{1}{n} \mathbf{W}^T \mathbf{H}(u_k) \widehat{\mathbf{D}}_X(u_k) \right] \\ &\quad \left[\frac{1}{n} \widehat{\mathbf{D}}_X^T(u_k) \mathbf{H}(u_k) \widehat{\mathbf{D}}_X(u_k) \right]^{-1} \sqrt{\frac{h}{n}} \widehat{\mathbf{D}}_X^T(u_k) \mathbf{H}(u_k) \boldsymbol{\varepsilon} \\ &= \frac{1}{\sqrt{h}} \sqrt{n}(\beta_0 - \bar{\beta})^T \frac{1}{n} \sum_{k=1}^n (\boldsymbol{\Psi}^T(u_k), 0) \\ &\quad \left(\begin{array}{cc} \Phi^{-1}(u_k) & 0 \\ 0 & \mu_2^{-1} \Phi^{-1}(u_k) \end{array} \right) \sqrt{\frac{h}{n}} \left(\begin{array}{cc} \Pi_X^T(u_k) & 0 \\ 0 & \Pi_X^T(u) \end{array} \right) \mathbf{D}_Z^T(u_k) \mathbf{H}(u_k) \boldsymbol{\varepsilon} + o_p(1/\sqrt{h}) \\ &= \frac{1}{n} (\beta_0 - \bar{\beta})^T \sum_{k=1}^n (\boldsymbol{\Psi}^T(u_k), 0) \\ &\quad \left(\begin{array}{cc} \Phi^{-1}(u_k) \Pi_X^T(u_k) & 0 \\ 0 & \mu_2^{-1} \Phi^{-1}(u_k) \Pi_X^T(u) \end{array} \right) \mathbf{D}_Z^T(u_k) \mathbf{H}(u_k) \boldsymbol{\varepsilon} + o_p(1/\sqrt{h}) \\ &= (\beta_0 - \bar{\beta})^T \frac{1}{n} \sum_{k=1}^n [\boldsymbol{\Psi}^T(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k), 0] \mathbf{D}_Z^T(u_k) \mathbf{H}(u_k) \boldsymbol{\varepsilon} + o_p(1/\sqrt{h}) \\ &= (\beta_0 - \bar{\beta})^T \frac{1}{n} \sum_{k=1}^n \boldsymbol{\Psi}^T(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) \sum_{i=1}^n Z_i \varepsilon_i K_h(u_i - u_k) + o_p(1/\sqrt{h}) \\ &= \sqrt{n} (\beta_0 - \bar{\beta})^T \\ &\quad \frac{1}{\sqrt{n}} \sum_{i=1}^n Z_i \varepsilon_i \left[\frac{1}{n} \sum_{k=1}^n \boldsymbol{\Psi}^T(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) K_h(u_k - u_i) \right] + o_p(1/\sqrt{h}) \\ &= O_p(1) + o_p(1/\sqrt{h}) = o_p(1/\sqrt{h}). \end{aligned}$$

Similarly, we can obtain that

$$\mathbf{T}_2 = o_p(1/\sqrt{h}).$$

Next, we focus on \mathbf{P} . We have

$$\begin{aligned}
\mathbf{P} &= \frac{1}{nh} \sum_{k=1}^n \sqrt{\frac{h}{n}} \boldsymbol{\varepsilon}^T \mathbf{H}(u_k) \widehat{\mathbf{D}}_X(u_k) \left[\frac{1}{n} \widehat{\mathbf{D}}_X^T(u_k) \mathbf{H}(u_k) \widehat{\mathbf{D}}_X(u_k) \right]^{-1} \sqrt{\frac{h}{n}} \widehat{\mathbf{D}}_X^T(u_k) \mathbf{H}(u_k) \boldsymbol{\varepsilon} \\
&= \frac{1}{nh} \sum_{k=1}^n \sqrt{\frac{h}{n}} \boldsymbol{\varepsilon}^T \mathbf{H}(u_k) \mathbf{D}_Z(u_k) \\
&\quad \left(\begin{array}{cc} \Pi_X(u_k) \boldsymbol{\Phi}^{-1}(u_k) \Pi_X^T(u_k) & 0 \\ 0 & \mu_2^{-1} \Pi_X(u_k) \boldsymbol{\Phi}^{-1}(u_k) \Pi_X^T(u_k) \end{array} \right) \sqrt{\frac{h}{n}} \mathbf{D}_Z^T(u) \mathbf{H}(u) \boldsymbol{\varepsilon} + O_p(c_n/h) \\
&= \frac{1}{n^2} \sum_{k=1}^n \sum_{i=1}^n \sum_{j=1}^n \left[K_h(u_i - u_k) K_h(u_j - u_k) \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_k) \boldsymbol{\Phi}^{-1}(u_k) \Pi_X^T(u_k) Z_j \right. \\
&\quad \left. + \mu_2^{-1} K_h(u_i - u_k) \frac{u_i - u_k}{h} K_h(u_j - u_k) \frac{u_j - u_k}{h} \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_k) \boldsymbol{\Phi}^{-1}(u_k) \Pi_X^T(u_k) Z_j \right] + o_p(1/\sqrt{h}) \\
&= \frac{1}{n^2} \sum_{k=1}^n \sum_{i=1}^n \left[K_h^2(u_i - u_k) \varepsilon_i^2 Z_i^T \Pi_X(u_k) \boldsymbol{\Phi}^{-1}(u_k) \Pi_X^T(u_k) Z_i \right. \\
&\quad \left. + \mu_2^{-1} K_h^2(u_i - u_k) \left(\frac{u_i - u_k}{h} \right)^2 \varepsilon_i^2 Z_i^T \Pi_X(u_k) \boldsymbol{\Phi}^{-1}(u_k) \Pi_X^T(u_k) Z_i \right] \\
&+ \frac{2}{n^2} \sum_{k=1}^n \sum_{i < j} \left[K_h(u_i - u_k) K_h(u_j - u_k) \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_k) \boldsymbol{\Phi}^{-1}(u_k) \Pi_X^T(u_k) Z_j \right. \\
&\quad \left. + \mu_2^{-1} K_h(u_i - u_k) \frac{u_i - u_k}{h} K_h(u_j - u_k) \frac{u_j - u_k}{h} \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_k) \boldsymbol{\Phi}^{-1}(u_k) \Pi_X^T(u_k) Z_j \right] + o_p(1/\sqrt{h}) \\
&= \mathbf{Q}_1 + \mathbf{Q}_2 + o_p(1/\sqrt{h}).
\end{aligned}$$

Define $\nu_0 = \int K^2(t) dt$ and $\nu_2 = \int t^2 K^2(t) dt$, then

$$\begin{aligned}
\mathbf{Q}_1 &= \frac{1}{n^2} \sum_{k=1}^n \sum_{i=1}^n \left[K_h^2(u_i - u_k) \varepsilon_i^2 Z_i^T \Pi_X(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_i \right. \\
&\quad \left. + \mu_2^{-1} K_h^2(u_i - u_k) \left(\frac{u_i - u_k}{h} \right)^2 \varepsilon_i^2 Z_i^T \Pi_X(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_i \right] \\
&= \frac{1}{nh} \sum_{k=1}^n \frac{1}{n} \sum_{i \neq k} \left\{ \frac{1}{h} K^2[(u_i - u_k)/h] \varepsilon_i^2 Z_i^T \Pi_X(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_i \right. \\
&\quad \left. + \mu_2^{-1} \frac{1}{h} K^2[(u_i - u_k)/h] \left(\frac{u_i - u_k}{h} \right)^2 \varepsilon_i^2 Z_i^T \Pi_X(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_i \right\} + O_p \left(\frac{1}{nh^2} \right) \\
&= \frac{1}{nh} \sum_{k=1}^n \left\{ \nu_0 E \left[\varepsilon_i^2 Z_i^T \Pi_X(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_i \mid u_i = u_k \right] f(u_k) \right. \\
&\quad \left. + \mu_2^{-1} \nu_2 E \left[\varepsilon_i^2 Z_i^T \Pi_X(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_i \mid u_i = u_k \right] f(u_k) \right\} + o_p(1/\sqrt{h}) \\
&= \frac{\nu_0 + \mu_2^{-1} \nu_2}{nh} \sum_{k=1}^n \text{Trace} \left\{ \Phi^{-1}(u_k) \Pi_X^T(u_k) E \left[Z_i \varepsilon_i^2 Z_i^T \mid u_i = u_k \right] \Pi_X(u_k) f(u_k) \right\} + o_p(1/\sqrt{h}) \\
&= \frac{\nu_0 + \mu_2^{-1} \nu_2}{h} E \left\{ \text{Trace} \left\{ \Phi^{-1}(u_k) \Lambda(u_k) \right\} \right\} + o_p(1/\sqrt{h}).
\end{aligned} \tag{A.14}$$

Now we consider \mathbf{Q}_2 . We have $E(\mathbf{Q}_2) = 0$ and

$$\begin{aligned}
\mathbf{Q}_2 &= \frac{2}{n^2} \sum_{i < j} \sum_{k \neq i, j} \left[K_h(u_k - u_i) K_h(u_k - u_j) \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_j \right. \\
&\quad \left. + \mu_2^{-1} K_h(u_k - u_i) \frac{u_k - u_i}{h} K_h(u_k - u_j) \frac{u_k - u_j}{h} \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_j \right] + O_p \left(\frac{1}{nh^2} \right) \\
&= \frac{2}{nh} \sum_{i < j} \frac{h}{n} \sum_{k \neq i, j} \left[K_h(u_k - u_i) K_h(u_k - u_j) \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_j \right. \\
&\quad \left. + \mu_2^{-1} K_h(u_k - u_i) \frac{u_k - u_i}{h} K_h(u_k - u_j) \frac{u_k - u_j}{h} \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_k) \Phi^{-1}(u_k) \Pi_X^T(u_k) Z_j \right] + o_p(1/\sqrt{h}) \\
&= \frac{2}{nh} \sum_{i < j} \left\{ \int K(t) K \left(\frac{u_i - u_j}{h} + t \right) dt f(u_i) \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_i) \Phi^{-1}(u_i) \Pi_X^T(u_i) Z_j [1 + O_p(c_n)] \right. \\
&\quad \left. + \mu_2^{-1} \int t \left(\frac{u_i - u_j}{h} + t \right) K(t) K \left(\frac{u_i - u_j}{h} + t \right) dt f(u_i) \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_i) \Phi^{-1}(u_i) \Pi_X^T(u_i) Z_j [1 + O_p(c_n)] \right\} \\
&\quad + o_p(1/\sqrt{h}) \\
&= \frac{2}{nh} \sum_{i < j} g \left(\frac{u_i - u_j}{h} \right) f(u_i) \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_i) \Phi^{-1}(u_i) \Pi_X^T(u_i) Z_j [1 + O_p(c_n)] + o_p(1/\sqrt{h}),
\end{aligned}$$

where

$$g(s) = \int K(t) K(s+t) dt + \mu_2^{-1} \int t(s+t) K(t) K(s+t) dt.$$

Consider the variance of \mathbf{Q}_2 , then

$$\begin{aligned}
& \text{Var} \left[\frac{2}{nh} \sum_{i < j} g \left(\frac{u_i - u_j}{h} \right) f(u_i) \varepsilon_i \varepsilon_j Z_i^T \Pi_X(u_i) \Phi^{-1}(u_i) \Pi_X^T(u_i) Z_j \right] \\
&= \frac{2(n-1)}{nh^2} E \left[g^2 \left(\frac{u_i - u_j}{h} \right) f^2(u_i) \varepsilon_i^2 \varepsilon_j^2 [Z_i^T \Pi_X(u_i) \Phi^{-1}(u_i) \Pi_X^T(u_i) Z_j]^2 \right] \\
&= \frac{2(n-1)}{nh^2} E \left[g^2 \left(\frac{u_i - u_j}{h} \right) f^2(u_i) \varepsilon_i^2 \varepsilon_j^2 Z_j^T \Pi_X(u_i) \Phi^{-1}(u_i) \Pi_X^T(u_i) Z_i Z_i^T \Pi_X(u_i) \Phi^{-1}(u_i) \Pi_X^T(u_i) Z_j \right] \\
&= \frac{2(n-1)}{nh^2} E \left[g^2 \left(\frac{u_i - u_j}{h} \right) \varepsilon_j^2 Z_j^T \Pi_X(u_i) \Phi^{-1}(u_i) \Pi_X^T(u_i) \right. \\
&\quad \left. E(Z_i \varepsilon_i^2 Z_i^T | u_i, u_j, \varepsilon_j, Z_j) \Pi_X(u_i) f^2(u_i) \Phi^{-1}(u_i) \Pi_X^T(u_i) Z_j \right] \\
&= \frac{2(n-1)}{nh^2} E \left[g^2 \left(\frac{u_i - u_j}{h} \right) \varepsilon_j^2 Z_j^T \Pi_X(u_i) \Phi^{-1}(u_i) \Lambda(u_i) \Phi^{-1}(u_i) \Pi_X^T(u_i) Z_j f(u_i) \right] \\
&= \frac{2(n-1)}{nh^2} h \left\{ \int g^2(t) dt E \left[\varepsilon_j^2 Z_j^T \Pi_X(u_j) \Phi^{-1}(u_j) \Lambda(u_j) \Phi^{-1}(u_j) \Pi_X^T(u_j) Z_j f^2(u_j) \right] + O_p(h) \right\} \\
&= \frac{2 \int g^2(t) dt}{h} E \left\{ \text{Trace} \left\{ \Phi^{-1}(u_j) \Lambda(u_j) \Phi^{-1}(u_j) \Pi_X^T(u_j) Z_j \varepsilon_j^2 Z_j^T \Pi_X(u_j) f^2(u_j) \right\} \right\} + O_p(1) \\
&= \frac{2 \int g^2(t) dt}{h} E \left\{ \text{Trace} \left\{ \Phi^{-1}(u_j) \Lambda(u_j) \Phi^{-1}(u_j) \Lambda(u_j) f(u_j) \right\} \right\} + O_p(1).
\end{aligned}$$

By the argument similar to Theorem 5 in Fan et al. (2001), we have

$$\frac{\mathbf{Q}_2}{\sqrt{\frac{2 \int g^2(t) dt}{h} E \left\{ \text{Trace} \left\{ \Phi^{-1}(u_j) \Lambda(u_j) \Phi^{-1}(u_j) \Lambda(u_j) f(u_j) \right\} \right\}}} \xrightarrow{d} N(0, 1). \quad (\text{A.15})$$

Finally, we have

$$\begin{aligned}
\lambda_n &= \frac{\text{RSS}_0 - \text{RSS}_1}{\text{RSS}_1/n} \\
&= \frac{\mathbf{Q}_1 + \mathbf{Q}_2 + \mathbf{T}_1 + \mathbf{T}_2}{\text{RSS}_1/n} + o_p(1/\sqrt{h}) \\
&= \frac{\mathbf{Q}_1 + \mathbf{Q}_2}{\text{RSS}_1/n} + o_p(1/\sqrt{h}).
\end{aligned} \quad (\text{A.16})$$

By Lemma 5 and (A.14), (A.15), (A.16), we complete the proof. \square

References

- Fan, J., Zhang, C., and Zhang, J. (2001), “Generalized likelihood ratio statistics and Wilks phenomenon,” *Annals of Statistics*, 29, 153–193.
- Mack, Y. P. and Silverman, B. W. (1982), “Weak and strong uniform consistency of kernel regression estimates,” *Probability Theory and Related Fields*, 61, 405–415.